Are Online Labor Markets Spot Markets for Tasks?:
A Field Experiment on the Behavioral Response to Wage Cuts

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Abstract

In some online labor markets, workers are paid by the task, choose what tasks to work on and have little or no interaction with their (usually anonymous) buyer/employer. These markets look like true spot markets for tasks rather than markets for employment. Despite appearances, we find via a field experiment that workers act more like parties to an employment contract: workers quickly form wage reference points and react negatively to proposed wage cuts by quitting. However, they can be mollified with “reasonable” justifications for why wages are being cut, highlighting the importance of fairness considerations in their decision-making. We find some evidence that “unreasonable” justifications for wage cuts reduce subsequent work quality. We also find that not explicitly presenting the worker with a decision about continuing to work eliminates quits, with no apparent reduction in work quality. One interpretation for this finding is that workers have a strong expectation that they are party to a quasi-employment relationship where terms are not changed and the default behavior is to continue working.

Keywords: Economics of IS; Electronic Commerce; Field Experiments; IT and new organizational forms

1 Introduction

According to Coase (1937), the boundary of the firm is determined by the relative costs and benefits of organizing production through markets or through authority. For the labor input to production, this markets-vs-authority choice was conceptualized by Simon (1951) as the choice between what Simon called a “sales contract” and the conventional employment contract. In Simon’s model, the employee accepts a wage in exchange for giving the employer control over what precise task, from some set of tasks, is done in the future. While employment is at will, the expectation is that the relationship will be on-going and persist until explicitly dissolved. In contrast, a sales contract is narrower in scope, with the worker agreeing to complete a specific task for a specific price. When the task is completed and payment made, the contract is dissolved. Simon’s labels are not always easy to apply (e.g., consider the salesperson
who works on commission), but the number of edge cases seems to have grown dramatically in recent years with the emergence of new labor-intermediating platforms. Platforms differ, but a number seem to create quasi-employment or quasi-sales contracts that share features of both types of relationships.1

One kind of new labor-intermediating platform of particular interest is the online labor market (Horton, 2010). In the case of online labor markets, some create relationships that at least look like traditional employment relationships: workers are paid by the hour, projects can be substantial and may require a number of different tasks, with the work being closely monitored and directed by the hiring firm. Examples from this mold include oDesk/Elance (now Upwork), Freelancer.com, Guru and several other more specialized sites. However, other online labor markets are far more task-focused and seemingly create sales contract relationships.

The most extreme version of a task-focused market is Amazon Mechanical Turk (MTurk), where tasks often take seconds and pay pennies. “Employers” in this market cannot exercise Simon-style authority because they do not communicate with workers—workers simply and immediately complete the task proposed by the buyer at the terms proposed by the buyer. Workers are free to take any task they want at the offered price. This market, at least based on its main characteristics, seems like a true spot market for tasks. If this is the case, MTurk and markets like it would offer firms a profoundly different way of obtaining labor inputs. With a spot market, firms could buy discrete chunks of labor from a global pool of workers at a market price, similar to how they obtain any other factor of production.

Even if MTurk is conceptualized by employers and Amazon itself as a spot market in tasks, this does not imply that workers share this view or behave accordingly. In this paper, we test whether workers on MTurk react as parties to a sales contract or an employment contract when presented with a wage cut. We conducted an experiment in which we contracted with workers for a data-entry task, paid them a high piece rate and then offered some treated workers the opportunity to keep working, albeit for a lower rate. While in both the sales contract and employment characterizations fewer workers should be willing to accept the lower offer, the two views differ in how “justifications” for the new offer should affect worker uptake.

1An example of a technology-driven platform creating this kind of confusion is Uber. Uber sets ride prices and determines who is eligible to drive, but does not tell drivers which precise customer to pick up or when to work—but it does impose standards of performance. Perhaps unsurprisingly, a number of worker classification lawsuits have sprung up around this industry.
For a party to a sales contract in a spot market, how price changes are justified is materially irrelevant and should not differentially affect uptake of the follow-on task. However, for a party to an employment contract, a lower offer would be perceived as a wage cut, which in turn could provoke retaliation from the worker, depending the perceived justice of that offer. We designed the different framings and justifications so that they would provoke workers to view the cuts as either reasonable and fair or capricious and unjust.²

Consistent with the employment view but not the spot market view, we find that workers were more likely to reject low offers, but justifications we deemed “reasonable” largely nullified this effect. Not all justifications were effective—justifying the cut in terms of our profits actually increased quits. This “profits” treatment also possibly reduced cooperation from workers. Unlike our quits and the consummate performance measures, we can only measure post wage cut quality for people who actually accepted the wage cut. It is possible that different kinds of workers quit in different treatments, so differences in quality across treatments could simply reflect selection effects as well as treatment effects. However, when we control for work quality from the first stage, we do find evidence of lower quality in the “profits” treatments, suggesting the possibility of retaliation.

In each of the different testing justifications for the new offer, we still explicitly framed the decision about whether to continue working as a choice, with the new per-task payment made salient. However, in one cell, we simply presented another data entry task with the new wage cut price clearly labeled, but without an explicit framing of a new offer and the need for a worker decision. In this cell, 100% of workers simply completed the next paragraph—a higher percentage than when they were offered their previous wage. Worker mistake is a possibility—perhaps they believed that not completing our extra paragraph would jeopardize payment—but we find no evidence of follow-on retaliation in this group. A parsimonious explanation could be that workers regard themselves as being in quasi-employment relationships and so the “default” is to simply keep working without the expectation of—or the need for—renegotiation.

The evidence from the experiment suggests that even in the most spot market-like of online labor markets, workers still bring an employee-like behavior to their interactions. MTurk is of course only one

²While we technically were offering a lower piece rate rather than a chance in time-based compensation, we refer throughout to the text as a “wage cut.”
market and the results of the experiment come from a self-selected sample of workers on that market. It is necessarily unclear whether these results would extend to other markets and other kinds of workers. However, given that MTurk is the most spot market-like in its observable characteristics, it seems probable that the employee-like behavior found on MTurk would generalize to other online labor markets.

2 Conceptual Framework

Developments in information and communications technology have made it possible for firms to obtain labor inputs from around the world. Given that labor markets have historically been strongly segmented by geography, it is unsurprising that an enormous literature sprang up around IT offshoring and outsourcing. This literature has focused on the decision to offshore and the sourcing mechanism (Tanriverdi et al., 2007); the structure (Chen and Bharadwaj, 2009); formality (Tanriverdi et al., 2007) and flexibility of contracts (Gopal and Koka, 2012); the interplay with the open source community (Ågerfalk and Fitzgerald, 2008); and so on. In this work, labor may cross borders, but is still being mediated by arm's length contracts between firms. However, the emergence of online labor markets lets firms buy labor inputs from workers directly, regardless of geography. These markets raise the question, however, whether firms are obtaining labor inputs as completed tasks, obtained in a spot market, or are they actually creating micro-employment relationships?

How can we distinguish a spot market for tasks from a market for employment? In this paper, we propose using the worker response to a proposed price change, i.e., a wage cut, as our distinguishing test. Our proposition is that if workers response to wage cuts in an employee-like way, then we have evidence against the spot market characterization. Of course, we must define what "employee-like" means. Fortunately, for this question, we can rely on a substantial empirical literature, primarily found in labor economics on the response to wage cuts.

A distinguishing feature of labor markets is that prices generally do not fall when demand falls (as in a recession). This fact has profound policy implications, as without falling wages, labor markets do not clear and some workers are involuntarily unemployed. The most prominent explanation for why firms do not cut wages is that workers with employment contracts—unlike say vendors with sales contracts—care a great deal about the fairness of wages and any proposed wage changes (Bewley, 1999). As such,
firms worry that wage cuts could be perceived as unfair and workers would retaliate explicitly, say by sab-
otaging the firm, or implicitly, by withholding their cooperation and “consummate performance.” What
is considered “unfair” often seems to depend strongly on what occurred in the relationship before the
cut: laboratory experiments support the proposition that past wage experiences can create wage refer-
ence points (e.g., Fehr et al. (2006)). It is the the workers’ expected reaction to the perceived unfairness
of wage cuts that makes them relatively rare. When wages have been cut, workers seem to respond as the
theory would predict. ³

It is theoretically unclear why fairness considerations matter so much more in employment trans-
actions compared to other kinds of commercial transactions. One possibility is that employment and
sales contracts fundamentally differ in the expectations parties have about when a contract has ended.
As a thought experiment, suppose we bought a single widget from a vendor for $1. If we returned the
second day and were told that a widget was now $1.01 (and the price was clearly posted for everyone as
such), we may not like the change but it would be difficult to argue the change was unfair or exploitative.
If in contrast, on our first day we agreed to a price of $1 but when we were rung up, we were told the
price had risen to $1.01, we would likely be rightfully angry at the unfairness of the new “offer.” The key
difference is that in the first scenario, after the initial widget was bought, the original contract was over
and both parties would expect an additional transaction to require a new contract. In the second sce-
nario, the contract was not yet completed and so the new offer was an illegitimate, exploitative attempt
to re-negotiate.

To bring us back to the labor context, an employer proposing a lower wage for the same set of tasks
is trying to re-negotiate the contract: the employment contract already specifies the payment if the em-
ployer wants more of some task performed—namely the price they already agreed upon. In the Simon

³In a real effort field experiment Kube et al. (forthcoming), cut wages, or more accurately, frustrated wage expectations, by
advertising the job using ambiguous language and then exploiting this ambiguity to pay some workers (college undergraduates)
less than they expected, causing substantial outputs in reduction. Although field experiments offer a more compelling test of
gift exchange than the laboratory experiments, observational data from real, long-term employment scenarios would be more
convincing, but the circumstances needed for causal inference—idiosyncratic factors leading to wage changes for one group of
workers but not for another—are rare. Occasionally, this kind of scenario occurs, and the evidence from them is consistent with
the view that negative reciprocity can be (but is not necessarily) long lasting and harmful to the organization. Mas (2006) found
that New Jersey police forces losing arbitration closed fewer cases for several months after the decision; Lee and Rupp (2007)
found that airline pilots subject to large, industry-wide wage cuts were late more often. In the airline example, the effects were
modest and transitory, perhaps because many airlines were at or near bankruptcy when the cuts were made, thus muting any
fairness judgments.
view, this ceding of discretion over tasks in exchange for a fixed wage is the essence of employment. In contrast, a buyer with a sales contract proposing a lower price for an additional unit of task is not trying to re-negotiate—once the task was completed, the original contract was completed.4

Unfortunately for us, a worker’s decision to accept a wage cut does not by itself distinguish the employment versus sales contract characterizations. The reason is that the newly offered wage might simply be below the quitting worker’s reservation wage. However, what does distinguish the two characterizations is the importance of context. When fairness concerns are motivating behavior, the context of the wage cuts matter, but should be immaterial to workers making decisions based solely on the costs and benefits.

Fairness has made a somewhat slow entry into economics, with general theories of fairness not appearing until the 1990s Fehr and Gächter (2000); Fehr and Schmidt (1999); Rabin (1993). As it is, those extant theories and the empirical underpinnings are still rather controversial (Binmore and Shaked, 2010). Despite disagreements, it is clear that individuals view some changes in price (and the resultant differences in allocations) as being more or less fair, depending on the context. For example, Kahneman et al. (1986), who surveyed people on their beliefs about a variety of economic factors found particularly strong opinions about labor market matters: Kahneman et al. found that people (a) view wage cuts that keep the firm solvent differently from cuts that increase already positive profits or exploit market changes, (b) see wage reference points as attached to specific workers in specific jobs, and not transferable to new employees or as having any relevance after sufficiently large reorganizations. The special importance of fairness in labor markets was emphasized by Rees (1993), who offers a number of examples of “real life” wage setting where fairness considerations were paramount. A common theme in all of examples from Rees and from Kahneman et al. is the importance placed on context and justification for a price change.

The research on organizational justice also highlights the importance of context. Colquitt et al. (2001), in a summarization of 25 years of research on organizational justice trace the early theories of what’s “fair” that looked only at inputs and outputs, to further elaborations that put emphasis on the importance of equality, need and the procedure by which some outcome came about. For example, to

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4This employment versus sales contract distinction is similar to the fundamental distinction Bajari and Tadelis (2001) draw between fixed-price and cost-plus procurement contracts, with the former analogous to the sale contract and the cost-plus contract analogous to the employment contract.
slightly paraphrasing an example from Rees (1993), an office worker learning he got a $2/hour raise would be quite pleased—but would be considerably less so if he learned every other worker in the office got a $3/hour raise (or even a $2.01 raise). Similarly, procedural justice matters—if same office worker learns that raises were based on strict seniority and he happened to be the most junior worker, he might view the situation very differently if he learned that raises went to the college friends of the HR director. In short, context matters, though as far as we are aware, there has been no unified theory that can map all of these different contexts to different popularly held fairness judgments.

Context can be manipulated, either by design or by chance, but the only study we are aware of that exploits variation in context is Greenberg (1990), who looked at employee theft in two factories following a temporary 15% pay cut. Employee theft rose in both plants, but in the plant where the CEO spent over an hour explaining the cuts and fielding questions, theft rose less than in the plant where management provided only a cursory explanation. The two plants were randomly assigned to treatment by the experimenter, though arguably the sample size was just two plants.

In our study, if context can reduce or exacerbate quits by workers, then the implication is that workers are behaving like employees. For our purposes, we do have some guides in which different contexts to try creating for our wage cuts (such as Kahneman et al. (1986)) but we are far away from having a tight connection between a theory of human fairness judgments and what contextual manipulations to “try.” We cut wages under different experimentally induced contexts and then give workers opportunities to respond by quitting, choosing a quality level if they accept and finally, the consummate performance they offer. Based on how workers react to those cuts, we can infer whether they perceive themselves as party to a sales contract—in which there is no expectation of continuation, contracts are complete and only offered wages matter—or party to an employment contract with reference points about wages, an expectation to continue working at the “old” wage and concern with the fairness of any offer.

It is ambiguous ex ante whether these fairness-mediated reactions would be more or less important

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5 We are not the first researchers to experimentally reduce wages—in two psychology studies from the early 1970s, one by Pritchard et al. (1972) and one by Valenzi and Andrews (1971), experimenters hired workers and randomly manipulated their wages or their perception of under- or over-payment. Andrews and Valenzi found that wage cuts, which the subjects knew were randomly determined, caused quits, but did not affect productivity. Pritchard et al. found that switching people between a wage and a quasi-piece-rate lowered productivity when a subject's past experience with the payment system led them to believe they were underpaid in the new system. Our study differs in that workers did not know wage cuts were random nor did workers switch between wages and piece-rates.
online. In our experimental setting, the short duration, anonymity and “one shot” nature of interactions could encourage the sales contract view. As interactions are one-shot, workers have less fear of being seen as an easy mark, as they might in a repeated interaction. Further, the limited time frame makes the formation of reference points presumably more difficult. The very low-stakes cut has an ambiguous effect: on one hand, the wage cuts are absolutely small and thus perhaps less offensive, yet the small stakes make principled refusals less costly. In conventional jobs, quitting has far greater consequences and we would expect workers to be less sanguine about losing their jobs, even after a wage cut. The higher stakes and longer duration of real jobs would probably make any fairness judgment more strongly felt. These more salient fairness judgments might magnify the response to wage cuts, though perhaps reactions with direct income consequences for the worker (such as quitting) are less likely and withdrawal of cooperation is more likely.

2.1 Why should we care about online labor markets?

To date, there has been relatively little work examining the nature of online labor markets. Most studies have focused on using them as a testing domain for some question of broader interest, such as (Pallais, 2014; Pallais and Sands, 2013). However, exceptions include Horton (2010); Kittur et al. (2013); Horton (2011). These markets are relatively new, which could explain the lack of research, but another factor is that they have not received much mainstream attention. Historically, total labor income in the United States have been about 2/3rds of GDP, or on the order of $8 trillion. By this standard, online labor markets are minuscule. Furthermore, MTurk is not the largest of online markets (for statistics on MTurk, see (Ipeirotis, 2010); for statistics on oDesk (now Upwork), see Agrawal et al. (2013)).

However, they are worthy of research attention for several reasons beyond just their current size and the ease with which research can be conducted (see Horton et al. (2011) for an elaboration on this argument). First, they are labor markets that are so comparatively simple that they can potentially give us insights into markets more generally: in the same way that studying gravity and classical mechanics is easier without friction, studying markets without geography, social networks, strong institutions and so on can be clarifying. Second, these markets are serving as a testing ground for novel combinations of human intelligence and machine intelligence—particularly in the case of MTurk. The ability to obtain
labor inputs algorithmically as part of a larger system is something genuinely new in the world and perhaps unsurprisingly, some of the heaviest users of MTurk are academic computer scientists. To make this point more concretely, in the recent proceedings of the “Human Computation” workshop, which is part of the AAAI (Association for the Advancement of Artificial Intelligence) annual conference—the premier AI conference—use of MTurk was dominant: of the 25 papers from the 2014 session, 23 were empirical, of which 16 used a marketplace for workers (as opposed to volunteers or users on social networking sites, for example). And of the 16 making using of online workers, all but 4 used Mechanical Turk—or 75%. Computer scientists are primarily building and testing new systems that take advantage of the “labor as a service” aspect of these markets. It seems plausible that this academic interest could eventually translate into widespread application and perhaps fundamentally change how labor is supplied and demanded.

2.2 What is different about online work?

The most obvious distinguishing characteristic of online work is that it happens online rather than by workers that are physically co-located. What is important about this is not likely to be the “online” face-to-face aspect per se; although a skype call or a Google hangout are not yet substitutes for face-to-face interactions, as communications technology improves, the qualitative differences between physical co-location and remote collaboration will presumably decline. What matters about online work is that it removes the role of geography. Geography profoundly shapes labor markets, in that throughout history, workers have needed to live relatively “close” to the productive capital and their fellow workers. This fact had implications not only for what was made where and by whom, but also the human capital decisions made by individuals and even the degree of specialization of workers. The differences in the returns to different skills based solely on geography—and the arbitrage opportunity this creates—has been vividly characterized as “trillion dollar bills on the sidewalk” Clemens (2011). In addition to these arbitrage opportunities, online labor markets change individual incentives to acquire skills. They might even allow for more specialization as the extent of the market grows—in an online labor market, a worker can reasonably specialize in some skill for which total global demand might only be 40 hours per week.

There are other distinctive features of online work. In contrast to conventional markets, it is com-
monplace for workers to work for many different “employers” within a short amount of time, even in 
markets like oDesk/Upwork that “look” more traditional. And as projects tend to be short-lived, there 
is far more individual variability in hours-worked. The shorter duration of projects makes the search 
and screening process both more important, as it must happen so much more frequently. At the same 
time, the short duration of projects makes the stakes for any particular hire that much lower. The case 
of MTurk—where there is no employer screening at all—is the exception that proves the rule, in that the 
stakes are so low for any “hire” that the default is for any worker to work on a task without approval from 
the employer. Unsurprisingly, all online labor markets have created features that try to make the search 
and matching process easier and faster: would-be firms can post jobs, get applicants, screen applicants 
and make a hire in a matter of hours.

In the case of MTurk, because workers search over the pool of tasks, employers need to attract work-
ers to their tasks if they hope to get them completed. Chilton et al. (2010) shows that employers delib-
erately add and remove their tasks from the marketplace at a high frequency so that their tasks appear 
higher in the search results shown to workers. This stands in sharp contrast to the “normal” way that 
firms get work done, which is to simply instruct their employees to perform the task. Firms competing 
to get workers to complete their tasks—rather than competing for employees—does suggest that online 
labor markets can at least look like true labor spot markets.

3  Empirical context, conditions for causal inference and the real-effort task 
employed

The experimental design was quite simple, even if the particular details of each cell—and the motiva-
tions for including that cell in that experiment—require more detailed explanation. We will first describe 
the experiment in broad strokes and then introduce the details of cells when we present the results. The 
experiment had three phases: (1) the wage expectation building phase, (2) the treatment phase and the 
(3) the “games” phase. In phase (1), workers performed a real effort task at a high piece rate. Then, in (2), 
they were assigned to different treatment groups and, in most cases, offered a new lower wage to per-
form an additional unit of the task. This wage cut was justified in different ways across cells. If workers
accepted the offered wage, they performed one more unit of the task. Finally, in (3), all subjects played a
series of contextualized games designed to measure their attitudes towards us as their employer/trading
partner. The goal of this phase was to measure any change in worker “consummate performance” caused
by the various treatments (Hart and Moore, 2008). These games revealed little and we confine the report-
ing of game outcomes to Appendix A.

Before discussing the experimental design in more depth, it is useful to explain the testing domain—
MTurk—and explain how the conditions for causal inference can be met in this domain. We will also
describe how we recruited subjects, the nature of the real-effort task and the games we used to try to
measure consummate performance.

3.1 Subject pool recruitment

We recruited experimental subjects from MTurk, an online labor market. Through an interface provided
by MTurk, registered users perform tasks (posted by buyers) for money. The tasks are generally simple
for humans to do yet difficult for computers—common tasks are captioning photographs, extracting data
from scanned documents and transcribing audio clips. Buyers control the features and contract terms
of the tasks they post: they choose the design, piece-rate, time allowed per task, how long each task will
be available and how many times they want a task completed.\footnote{Tasks are often done multiple times by different workers for quality-control purposes.}

Workers, who are identified to buyers only by a unique string of letters and numbers, can inspect
tasks and the offered terms before deciding whether to complete them. Buyers can require workers to
have certain qualifications, but the default is that workers can “accept” a task immediately and begin
work. Once the workers submit their work, buyers can approve or reject their submission. If the buyer
approves, MTurk pays the worker with buyer-provided escrow funds; if the buyer rejects, the worker is
paid nothing.

Although most buyers post tasks directly to MTurk, it is possible to host tasks on an external site that
workers reach by following a link. We used this external hosting method; we posted a single placeholder
task containing a description of the work and a link to follow if subjects wanted to participate. When they
completed all stages of the experiment on the external site, they received a completion code which they
entered into the original MTurk interface. Figure 4 shows the landing page that all subjects first arrived at, regardless of assignment.

### 3.2 Real-effort task

For the real-effort task, subjects transcribed (not translated) paragraph-sized chunks of Adam Smith’s *The Wealth of Nations*. A sample paragraph—and the interface subjects used to transcribe the paragraph—is shown in Figure 5 in Appendix B. This task was sufficiently tedious that no one was likely to do it “for fun” and it was sufficiently simple that all market participants could do the task. The source text was machine translated into Dutch and blurred, which increased the error rate of the transcriptions, thereby providing a more informative measure of work quality. Translating the text also prevented subjects from finding the text elsewhere on the Internet. Although this task might seem unusual and raise suspicions that the workers were in some kind of experiment, we view this as unlikely. Transcribing text from images is a commonplace use of MTurk. In fact, it is one of the few tasks that Amazon now supports with pre-made templates and pools of vetted workers (though this was not the case when our experiment was run). In Appendix B, Figure 12 shows the template and shows other, similar commonplace MTurk tasks.

### 3.3 Conditions for causal inference

To identify causal effects from the treatments, we needed ways to (a) as-good-as randomly assign subjects to different groups, (b) keep subjects from changing groups once assigned (c) keep subjects from participating multiple times, and (d) minimize non-random attrition. These problems are somewhat more challenging in an online setting; *Horton et al.* (2011) discusses these challenges and explains how they can be addressed, as well as discussing the many advantages of using online populations—particularly those in labor markets—as experimental subjects. *Steelman et al.* (2014) also highlights the usefulness of sampling from non-traditional populations accessed online.

For (a): we sequentially assigned subjects to treatment groups as they accepted our task, stratifying on arrival order. For example, arrival 1 went to the first cell, arrival 2 to the second cell and so on and so forth.

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7Since the text was presented as images, subjects were unable to simply copy and paste the text into the text box on the form. It is irrelevant whether or not any subjects actually knew Dutch since, if anything, knowledge of the language might make the task more difficult given the poor quality of the translation. One subject that apparently did speak Dutch sent an email warning us that the work was “grammatical gibberish.”
when all cells were exhausted, the next subject would go to the first cell again. Unbeknownst to subjects, clicking on our link initiated the experimental group assignment. As subjects neither knew nor could control their relative arrival order, this assignment mechanism meets the unconfounded assignment condition necessary for estimating causal effects. In fact, this method provides more precise estimates of causal effects. The reason for the improved precision is that stratification on arrival order improves balance on arrival times, which in turn are correlated with demographics (because of time zones) and behaviors (early arrivals are more likely to be heavy users of the site). Our method for achieving balanced samples was effective: for the binary covariates of gender, resident in the U.S. or Canada and whether the worker spends more than 10 hours a week online doing tasks for money, none of the p-values for a χ² test are conventionally significant. The actual demographic survey used to collect these measures is shown in Appendix B, Figure 11. Note that despite stratification, because of attrition during the pre-randomization phase, the realized sample size for each experimental group differs by more than one.\(^8\)

For (b): if workers were aware of the different treatment groups, they would have an incentive to get into a group with a larger payoff. Even though subjects were unaware of the other treatments, to prevent the possibility of subjects hunting for the “best” treatment group, the software assigning subjects to groups tracked users’ IP addresses. This tracking prevented subjects from changing their assignment after their initial “assignment” click.

For (c): workers with multiple online identities could in principle complete our task multiple times. However, this double-dipping is unlikely: because buyers often want the same task done multiple times by different workers, MTurk designed several policies and software features to prevent a worker from having more than one account.\(^9\)

For (d): to prevent differential attrition driven by differences among treatments, all subjects had identical initial experiences during the wage expectation building phase. Subjects’ experiences differed by group only after already performing three transcriptions. Because of this investment, there was no attrition after subjects observed their own treatment. An important point is that although we assigned

\(^8\)If we had no attrition, then stratification should lead to sample sizes differing by at most one by experimental group.

\(^9\)Workers must agree to have only one account—any detected attempt to have multiple accounts leads to a permanent ban. MTurk also requires browsers to accept cookies. If a person had two bank accounts or credit cards and two separate computers not sharing a network connection, they could in principle participate twice, but given the low stakes and risks involved (if this was detected, they would be banned from the site), we consider this possibility highly unlikely.
subjects to treatments at the moment they accepted the task, the conceptual point of randomization was after the wage expectation stage, but before the treatment stage.

The pre-randomization attrition is easy to spot in the data: attriting subjects just stop working before the end of the paragraph, never get to the treatment stage and never submit their completion code to receive payment. We dropped from the sample any worker making more than 100 errors on a paragraph. This occurred for a little less than 20% of all subjects who started the paragraphs, but since these individuals were dropped prior to conceptual randomization, this kind of attrition is immaterial.

3.4 Measuring outcomes following the wage cut

We can easily measure whether a worker was willing to transcribe a fourth paragraph following a wage cut. Measuring the quality of their work is somewhat more challenging. To measure work quality, we calculated the minimum number of “edits” would be required to transform the provided transcription to a perfect transcription of the underlying text. For example, if the worker typed “infro

information “ instead of “information “, we could fix the error with one one simple transposition edit, namely by switching the “r” and the “o.” Note that this would be fewer edits than the two edit adjustment of changing the “r” to an “o” and the “o” to an “r.” This minimum-number-of-edits metric is called the edit distance, or the Levenshtein distance.

In addition to output and work quality, we also had subjects play a series of contextualized games with us to measure their “consummate performance” on the task. The design of these games and the results from the experiment are Appendix A.

4 Experiment design and methodology

Figure 1 shows how subjects flowed through the experiment, from the expectation-building transcription phase, to the treatment phase and finally, the “games” phase. In the wage expectation building phase, subjects transcribed three paragraphs at a rate of 10 cents per paragraph.10 If subjects finished three paragraphs in about 100 seconds, so the offered payment of 10 cents per paragraph is equivalent to $3.60/hour ($28.80 per day). It is challenging to assess wages on MTurk, as prices are for tasks, not for time, but evidence suggests that the wage we paid is considerably higher than the median reservation wage on the platform. Horton and Chilton (2010) estimates that reservation wages are approximately log normally distributed and centered at $1.38.
paragraphs, they moved to the treatment phase. In the treatment phase, subject experiences differed based on group assignment. Generally, subjects were offered some new piece rate to transcribe a fourth paragraph. When a lower wage was proposed, it was always 3 cents—a 70% reduction from their previous wage.

Subjects assigned to Control \((n = 23)\) were given the option to transcribe another paragraph at their previous 10 cent piece rate. Figure 2 shows the interface used for Control and lists the exact language used for the group where subjects were asked if they wanted to transcribe an additional paragraph. Screenshots of all experimental interfaces and instructions are in Appendix B.

Subjects in WageCut \((n = 17)\) were offered the option to transcribe another paragraph, but at a rate of only 3 cents. In FairnessPrimedWC \((n = 24)\), subjects were first asked what they considered a fair wage for the task, before receiving the 3 cent offer. Subjects in OnlyGames \((n = 21)\) were not given the option to transcribe an additional paragraph, but instead went straight to the games phase. Subjects in groups ProductivityWC \((n = 25)\), ProfitWC \((n = 23)\), and NeighborWC \((n = 25)\) were given the option to continue working at the lower rate of 3 cents, but the offer was justified differently across groups. We also had an UnframedWC \((n = 26)\) group in which workers were simply brought to fourth paragraph, without an explicit framing of a decision to keep or stop working, but with the fourth paragraph clearly labeled as only paying 3 cents. Following the treatment phase, regardless of whether they accepted the wage cut, all workers continued on to the games phase.

5 Results

5.1 Accepting an unexplained wage cut

We first test whether being offered a lower wage caused fewer workers to accept our offer to transcribe a fourth paragraph. At a lower offered piece rate, we expected fewer workers to accept the offer, since for some fraction, the marginal cost of doing another transcription outweighs the marginal benefit. This test cannot distinguish between the sales contract and employment views, but it does let us verify that workers find the task costly and that precise amount offered matters to workers. The two relevant experimental cells for this question are Control and WageCut.
Figure 1: Experimental design

Notes: This figure illustrates the flow of subjects through the experiment.

**CONTROL:** Subjects were offered their previous high wage of 10 cents to perform an additional transcription.

**WAGECUT:** Subjects were offered a lower wage of 3 cents to perform an additional transcription. No explanation was given for the lower wage offer. The precise language of the offer was “Will you transcribe an additional paragraph for 3 cents?“

Table 1 shows that, as expected, substantially fewer workers are willing to transcribe a paragraph when the offered rate is 3 cents instead of 10 cents.

Figure 2: Offer to perform an additional transcription presented to subjects in **CONTROL**

Notes: After subjects completed the initial three paragraphs, they were, in most cases, presented the offer to do an additional page. This is a screenshot for offer presented to subjects assigned to **CONTROL**.
Table 1: Effect of a wage cut on willingness of the worker to perform an additional task

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>Transcribe additional paragraph? (ADDPara = 1)</th>
</tr>
</thead>
<tbody>
<tr>
<td>(1)</td>
<td>(2)</td>
</tr>
<tr>
<td>Offered 3 cents for fourth transcription, WAGECut</td>
<td>−0.327**, (0.152)</td>
</tr>
<tr>
<td>From US/CAN</td>
<td>0.179, (0.192)</td>
</tr>
<tr>
<td>Male</td>
<td>−0.296, (0.185)</td>
</tr>
<tr>
<td>Prior Error</td>
<td>−0.234, (0.162)</td>
</tr>
<tr>
<td>Constant</td>
<td>0.739***, (0.099)</td>
</tr>
</tbody>
</table>

Comparison Group

<table>
<thead>
<tr>
<th>Observations</th>
<th>CONTROL</th>
<th>CONTROL</th>
</tr>
</thead>
<tbody>
<tr>
<td>40</td>
<td>40</td>
<td></td>
</tr>
<tr>
<td>R²</td>
<td>0.109</td>
<td>0.282</td>
</tr>
</tbody>
</table>

Notes: This table reports two OLS regressions where the dependent variable is an indicator for whether the subject was willing to transcribe a fourth paragraph. In the control (the omitted group), CONTROL, the offer was 10 cents—the same piece rate as for the previous work. For the treated group, WAGECut, the offer was 3 cents. In Column (2) regressors are added for several pre-treatment variables, including the indicators for whether the subject was from the US or Canada, a self-reported male, and a continuous variable which is the log cumulative prior errors on the three previous paragraphs. Robust standard errors are reported. Significance indicators: $p \leq 0.10 : *, p \leq 0.05 : **$ and $p \leq .01 : ***$.

Column (1) of Table 1 reports an estimate of

$$ADDPara = \beta_0 + \beta_1 WAGECut + \epsilon,$$

where ADDPara is an indicator for whether the worker performed the additional paragraph transcription and WAGECut is an indicator that the subject was assigned to that group (we will use this indicator/group name convention throughout the paper). The comparison group in the regression is CONTROL. In terms of magnitude, we can see that a 70% reduction in wages (10 cents to 3 cents) led to an approximately 30% reduction output, implying an elasticity of labor supply of about 0.42.

In Column (2), several pre-treatment subject characteristics are added to the regression model. With these added covariates, the treatment effect increases slightly relative to Column (1), but this estimate is well within the 95% confidence interval for the Column (1) treatment effect, as we would expect for a true experiment. There is some evidence that self-reported males were less likely to do the follow-on task, as were those that made a large number of prior errors during the expectation building phase. However,
none of the coefficients are conventionally significant. We can test for whether there are heterogeneous
treatment effects by comparing the Column (2) regression model with one in which the treatment indi-
cator is interacted with all of the pre-treatment covariates. When we perform this test, we fail to reject the
null hypothesis that the coefficients on the interaction terms are all zero, \( p = 0.414 \), making substantial
treatment effect heterogeneity unlikely, at least within the limits of the statistical power available.

5.2 Is offer rejection about fairness?

That fewer workers accept the offer to do an additional task is unsurprising—it is in a sense, over-
determined, as it is predicted by both the employment and sale contract views. However, we can partially
test whether workers have an “employment-like” view that a lower wage offer is a contract breach by (1)
getting subjects to think about fairness with respect to the task and then (b) presenting them the low
wage offer. Some subjects were assigned to a cell FairnessPrimedWC, in which they were first asked
what would be a fair rate for the task before receiving the 3 cent offer. Appendix B, Figure 6 shows the
interface used to present the “fairness” question to users.

**FairnessPrimedWC:** Subjects were offered a lower wage of 3 cents to perform an additional task. No
explanation was provided, but we first asked subjects what they thought would be a fair wage for
doing an additional task.

For the sake of brevity, we simply describe the results rather than present them in regression tables.
Comparing acceptances in FairnessPrimedWC to our WageCut cells, acceptance of the offer was 41% in WageCut and 33% in FairnessPrimedWC, or an 8 percentage point difference. Despite this lower
uptake when fairness was primed, the standard error for the differences in means is slightly more than
15 percentage points and the p-value of a t-test for a two-sided means comparison is \( p = 0.63 \). Clearly we
cannot reject a null hypothesis of no effect from priming, but the direction of the effect is what we would
expect if worker decision-making was mediated by fairness concerns (assuming a 3 cent offer after a 10
cent offer would be viewed by most as unfair).
5.3 The effects of justifying the wage cut

Providing some justification for a wage cut might change a worker's probability of accepting that wage cut. For example, a “reasonable” justification might induce the worker to simply consider the offer relative to their reservation wage for the task. In contrast, “unreasonable” justifications might even increase the cost of accepting the lower offer. To return to the question of how we should characterize relationships in this market, in the sales contract view, justifications should be immaterial because they do not change payoffs, but they would matter in the employment view. To test this notion that justifications can alter wage cut acceptance, we randomized subjects to cells with varying justifications.

**PRODUCTIVITYWC:** Subjects were offered a lower wage of 3 cents to perform an additional task, but we first informed them that most workers can complete tasks more quickly after they gain some experience. The precise language was “We find that workers do this task much faster after doing three paragraphs, would you be willing to do an additional paragraph for 3 cents?”

**PROFITWC:** Subjects were offered a lower wage of 3 cents to perform an additional task, but we first informed them that we are trying to get as much work done for as little money as possible. The precise language was “We are trying to get this work done for as little money as possible, would you be willing to do an additional paragraph for 3 cents?”

**NEIGHBORWC:** Subjects were offered a lower wage of 3 cents to perform an additional task, but we informed them that in our experience, other workers are willing to accept the lower offer. The precise language was “We find that some workers are willing to do a fourth paragraph for 3 cents. Would you be willing to do an additional paragraph for 3 cents?”

The motivations for **PRODUCTIVITYWC** and **PROFITWC** were to contrast “reasonable” and “unreasonable” justifications for the wage cut. Although **PROFITWC** is unrealistic in the sense that no real firm would propose a wage cut in that manner, it is still a useful treatment in that it lets us test how workers react to an extreme, insulting justification: **PROFITWC** tests whether any justification—no matter how insulting—“works.” Perhaps any justification would be effective—Langer et al. (1978) showed that

---

11: This was a true statement. From earlier pilot experiments, we knew that some subjects were willing to work at this lower wage.
asking someone to let you cut in line to make copies “because you need to make copies” dramatically increases compliance compared to just asking without the justification.

NEIGHBORWC tests the notion that workers “learn” how to react from other workers. One reason why NEIGHBORWC might mitigate the behavioral response to a wage cut is that perhaps the psychic cost of taking work below one’s past wage comes not just from a fear of being exploited, but also from a fear of being singularly exploited. The notion that social comparisons could affect the response to wage cuts has some empirical support—Luttmer (2005) found that individuals feel worse-off when others around them earn more.

Consistent with the employment view but not the sale contract view, the empirical results in Table 2 show that justifications do matter, with “reasonable” justifications substantially increasing the fraction of workers willing to work at the new, lower piece rate. Compared to the unexplained cut in WAGECUT, however, the unreasonable explanation, PROFITWC, did not substantially decrease acceptance, at least within the limits of available statistical power.

Table 2: The effects of wage cut justifications on worker output choice following a wage cut

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>Transcribe additional paragraph?</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
</tr>
<tr>
<td>Neighbor justification, NEIGHBORWC</td>
<td>0.228 (0.152)</td>
</tr>
<tr>
<td>Productivity justification, PRODUCTIVITYWC</td>
<td>0.268* (0.152)</td>
</tr>
<tr>
<td>Profit justification, PROFITWC</td>
<td>-0.107 (0.155)</td>
</tr>
<tr>
<td>“Reasonable” justifications (NEIGHBORWC + PRODUCTIVITYWC)</td>
<td>0.310*** (0.102)</td>
</tr>
<tr>
<td>Constant</td>
<td>0.412*** (0.118)</td>
</tr>
<tr>
<td>(2)</td>
<td></td>
</tr>
<tr>
<td>Comparison Group</td>
<td>WAGECUT</td>
</tr>
<tr>
<td>Observations</td>
<td>90</td>
</tr>
<tr>
<td>R²</td>
<td>0.101</td>
</tr>
</tbody>
</table>

Notes: This table reports two OLS regressions where the dependent variable is an indicator for whether the worker subject was willing to transcribe a fourth paragraph. Comparisons are being made against WAGECUT, in which the offer was 3 cents and no justification was offered for the cut. In Column (1), indicators for each justification—productivity increases, employer profits and comparison to “neighbor” choices are used as regressors. In Column (2), the “unreasonable” profit justification is pooled with the unexplained cut and the two “reasonable” justifications from G4 and G6 are pooled as an indicator for “Reasonable” justifications. Standard errors are robust. Significance indicators: $p \leq 0.10 : *, p \leq 0.05 : **$ and $p \leq .01 : ***$. 

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We first consider how each justification by itself affected a worker’s willingness to perform the fourth transcription at the new rate. Column (1) of Table 2 reports an estimate of

$$\text{ADDPara} = \beta_0 + \beta_1 \text{PRODUCTIVITYWC} + \beta_2 \text{PROFITWC} + \beta_3 \text{NEIGHBORWC} + \epsilon$$ (2)

where the sample consists of all the justification groups and the unexplained group, with \text{WAGECUT} being the comparison group. The coefficient on the productivity justification indicator is positive and conventionally significant, with nearly a 30 percentage point increase in the fraction of workers accepting the lower offer. The coefficient on the neighbor justification is smaller, with the effect being about 23 percentage points. It is not conventionally significant, though the difference between the neighbor and the productivity coefficients would itself not be significant. The profit justification is negative and non-trivial in magnitude—the effect is about an 11 percentage point reduction in task acceptance—but it is far from conventionally significant. In Column (2), the outcome is the same, but we pooled what we considered “reasonable” justifications, \text{PRODUCTIVITYWC} and \text{NEIGHBORWC}, and compared them to what we considered unreasonable or missing justifications, \text{WAGECUT} and \text{PROFITWC}. Here the difference is stark, with the reasonable justifications increasing output nearly 30 percentage points. This greater effect is attributable to the roughly similar magnitudes for the coefficients on \text{NEIGHBORWC} and \text{PRODUCTIVITYWC} and the negative and fairly large coefficient on the profit justification indicator.

### 5.4 The effects of explicitly asking about task continuation

One of the key differences between the sales and employment contract is that in the employment contract, the expectation is that the project will continue until explicitly dissolved. In contrast, the sales contract is ended when the specified task is completed. We wanted to test whether simply presenting workers the additional transcription task, labeled with the new 3 cent wage, would cause a different reaction from workers, relative to explicitly flagging the proposed new wage and requiring workers to give a yes or no answer to whether they wanted to continue. To test whether this offer framing mattered, we created a cell \text{UNFRAMEDWC}.

\text{UNFRAMEDWC}: Subjects were not explicitly offered a lower wage to perform an additional task. Rather,
they were simply brought to the next task, which was clearly labeled with the new, lower wage of 3 cents. There was a button at the bottom of the screen that allowed them to quit and continue to the follow-on contextualized games.

Consistent with the employment contract view, Table 3 shows that this “unframed” wage cut was remarkably effective, in that all subjects assigned to UNFRAMEDWC transcribed the fourth paragraph.

Table 3: The effect of not explicitly framing the wage cut as a decision to continue working

<table>
<thead>
<tr>
<th>Dependent variable: Transcribe additional paragraph? (ADDPARA = 1)</th>
<th>(1)</th>
<th>(2)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Unframed wage cut, UNFRAMEDWC</td>
<td>0.588***</td>
<td>0.261***</td>
</tr>
<tr>
<td>(0.101)</td>
<td>(0.090)</td>
<td></td>
</tr>
<tr>
<td>Constant</td>
<td>0.412***</td>
<td>0.739***</td>
</tr>
<tr>
<td>(0.078)</td>
<td>(0.065)</td>
<td></td>
</tr>
<tr>
<td>Comparison Group WAGECUT</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Observations</td>
<td>42</td>
<td>48</td>
</tr>
<tr>
<td>R²</td>
<td>0.460</td>
<td>0.155</td>
</tr>
</tbody>
</table>

Notes: This table reports OLS regressions where the dependent variable is an indicator for whether the worker subject was willing to transcribe a fourth paragraph. In Column (1), the independent variable is an indicator for assignment to UNFRAMEDWC, the unframed wage cut. The comparison group is WAGECUT. In Column (2) the comparison group is instead CONTROL, in which subjects received a 10 cent offer to perform an additional transcription. Standard errors are robust. Significance indicators: \( p \leq 0.10 : \ast, p \leq 0.05 : \ast \ast \) and \( p \leq 0.01 : \ast \ast \ast \).

First we consider output in the unframed cell relative to the explicit, unjustified wage cut of WAGECUT. Column (1) of Table 3 reports an estimate of

\[
ADDPARA = \beta_0 + \beta_1 \text{UNFRAMEDWC} + \epsilon.
\]  

(3)

The comparison group is WAGECUT, the unexplained wage cut. The coefficient on the “unframed” indicator shows that not explicitly framing work continuation as a choice leads to very high acceptance of the wage cut: in fact, 100% of subjects in UNFRAMEDWC performed the follow-on paragraph transcription, compared to just a little more than 40% of subjects in WAGECUT. Column (2) reports the same regression as Column (1), but with the comparison group being the CONTROL group instead of WAGECUT. When we compare the unframed wage cut group to the subjects that had no wage cut at all, we can see that the unframed group had about a 25 percentage point increase over those not receiving a wage cut at all. In
other words, not framing a decision to continue working led to far more output than even offering the previous rate.

There are several possible interpretations of these results. It is possible that subjects failed to consider their option of refusing to do the task by simply entering a poor transcription (the most extreme form being leaving the text box blank). Another possibility is a worker might believe they were in error: if a worker thought she had read the instructions incorrectly, she might worry that a skipped transcription would jeopardize payment for work already completed. These caveats aside, the evidence does suggest that explicitly flagging a wage change has a different effect than imposing the same change surreptitiously. If workers regard their relationship with the buyer as an employment contract, this continuation of work—without any explicit discussion of prices or decisions—is what they would expect to happen normally. In contrast, a worker with the sales contract view who was told ex ante they would perform three paragraphs would be very surprised by an additional task, as the contract was at that point dissolved.

5.5 Comparison of output by offer size, justification and explicitness of follow-on offer

The pattern of output results can be more clearly seen in Figure 3. We have divided reasons into “reasonable” and “non-reasonable” with “reasonable” based on our own judgment. The plots show that reasonable justifications keep output levels quite close to those in the 10 cent offer CONTROL group. The remarkable uptake in UNFRAMEDWC is also apparent, as well as the much lower output in the cells with no justification or an unreasonable justification.

5.6 Follow-on task performance

Workers might be willing to “accept” a wage cut and then retaliate by offering poor performance on the task—say by introducing many errors in transcription. This is the consensus view for why firms are so averse to cutting wages. In our experimental setting, detecting retaliation is challenging because we know that the different cells have different task acceptance rates. As such, any differences in follow-on error rate could reflect some combination of selection and treatment. We can potentially deal with this selection problem by controlling for transcription quality from the first phase of the experiment. It is
the case that error rates in the fourth paragraph are highly correlated with the error rate in the initial three paragraphs: when pooled, the point estimate of the correlation coefficient is $\rho = 0.33$ and a 95% confidence interval of $[0.14, 0.56]$.\footnote{Standard errors for the correlation coefficient were calculated with one million bootstrap replications.}

Table 4 shows that there is no evidence that workers willing to accept the wage cut were any worse overall, relative to Control. However, there is some evidence that workers assigned to the unreasonable ProfitWC produced worse transcriptions, even when controlling for output quality during the first phase of the experiment. However, it is important to remember that even with our prior error controls, we cannot credibly estimate causal effects.

First, we test whether workers assigned any of the wage cut cells had worse performance compared to those who were offered their previous rate. Column (1) of Table 4 reports an estimate of

$$\log\text{EDITDIST}_4 = \beta_0 + \beta_1\text{ANYWAGECUT} + \log \sum_{k \in \{1,2,3\}} \text{EDITDIST}_k + \epsilon, \quad (4)$$

where EDITDIST$^4$ is the edit distance for the fourth paragraph and ANYWAGECUT is an indicator that the subject was assigned to any of the wage cuts groups. Note that we control for the log cumulative error in the first three paragraphs, which should at least partially address the possibility that selection effects drive differences in error rates. The coefficient on ANYWAGECUT is negative—implying fewer errors—
Table 4: Transcription errors in fourth paragraph among workers willing to accept the new offer

<table>
<thead>
<tr>
<th></th>
<th>Dependent variable:</th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Log edit distance for the fourth paragraph</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
</tr>
<tr>
<td>AnyWageCut</td>
<td>−0.060</td>
<td>(0.098)</td>
</tr>
<tr>
<td>FairnessPrimedWC</td>
<td>0.020</td>
<td>(0.192)</td>
</tr>
<tr>
<td>NeighborWC</td>
<td>0.068</td>
<td>(0.170)</td>
</tr>
<tr>
<td>ProductivityWC</td>
<td>−0.007</td>
<td>(0.167)</td>
</tr>
<tr>
<td>ProfitWC</td>
<td>0.491**</td>
<td>(0.206)</td>
</tr>
<tr>
<td>UnframedWC</td>
<td>0.087</td>
<td>(0.159)</td>
</tr>
<tr>
<td>Prior Error</td>
<td>0.290***</td>
<td>0.264***</td>
</tr>
<tr>
<td></td>
<td>(0.082)</td>
<td>(0.092)</td>
</tr>
<tr>
<td>Constant</td>
<td>1.540***</td>
<td>1.500***</td>
</tr>
<tr>
<td></td>
<td>(0.319)</td>
<td>(0.375)</td>
</tr>
</tbody>
</table>

<table>
<thead>
<tr>
<th>Comparison Group</th>
<th>Control</th>
<th>WageCut</th>
</tr>
</thead>
<tbody>
<tr>
<td>Observations</td>
<td>96</td>
<td>79</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.117</td>
<td>0.168</td>
</tr>
</tbody>
</table>

Notes: This table reports OLS regressions where the dependent variable is the log edit distance for the fourth paragraph. In Column (1), the independent variable is an indicator for assignment to any group where wages were cut. The comparison group is Control. In Column (2) the comparison group is instead WageCut and regressors are indicators for the remaining cells in which wages were cut. Both regressions include controls for worker error in the first phase of the experiment. This prior error control is the log of the sum of edit distances for the first three paragraphs transcribed. Standard errors are robust. Significance indicators: $p \leq 0.10: \ast, p \leq 0.05: \ast\ast$ and $p \leq 0.01: \ast\ast\ast$. 
but is also imprecisely estimated and far from conventionally significant.

In Column (2), we decompose ANYWAGECUT into the indicators for the different cells and use WAGE-CUT as the comparison group. This lets us test whether any of the framings or justifications among workers accepting the wage cut affected output. Here we see that the coefficient on PROFITWC is large and positively significant, with nearly a 50% higher error rate. This is suggestive evidence that workers in the profits justification retaliated, but there are two important caveats. First, this group had the lowest acceptance rate and so any selection effects would be particularly strong for this group. Second, we are looking at five different treatment cells and only one is conventionally significant. On the other hand, for theory reasons, we also have reason to believe that if there was retaliation, it would be more likely in PROFITWC.

6 Conclusion

Our main positive finding is that framing can strongly affect the behavioral response to a wage change. Workers reduce output—and possibly work quality and cooperation—when wages are cut for capricious or selfish-seeming reasons, but workers do not reduce output as much and apparently are still willing to cooperate if the employer has seemingly valid reasons for his or her actions. Workers on MTurk exhibit the same fairness-mediated reactions to wage cuts thought to characterize worker reactions to wage cuts in conventional employment relationships. Further, the results from the “unframed” wage cut experiment suggest that workers have a bias towards simply continuing to work rather than viewing the completion of the last agreed-upon task as the end of the contract. While this is not the only interpretation, it is some evidence towards the employment characterization of the relationships that are created. As such, it seems unlikely that firms hiring in online labor markets can treat their interactions as spot market transactions but must instead consider the quasi-employment nature of the relationships they seem to create.

References

Ågerfalk, Pär J and Brian Fitzgerald, “Outsourcing to an Unknown Workforce: Exploring Opensurcing


A Consummate performance measures

When we designed the experiment, we also wanted to look for other indications of work effort and cooperation beyond just acceptance of the wage cut. To that end, we had subjects play a number of contextualized games to measure “consummate performance.” This part of the experiment had generally uninteresting results—in part because the relatively low power of the design. We do present the results here in an appendix for the benefit of interested parties.

A.1 Consummate performane measures

To measure consummate performance, subjects played a collection of contextualized games that we hoped would not seem unusual to the subjects. These games tried to measure worker patience, trust in us, and cooperation with us. To measure patience, we simply asked subjects whether they would be willing to forgo nearly immediate payment in exchange for a 30 cent bonus to be paid out in two weeks.

We measured trust with a modified version of the trust game (Glaeser et al., 2000). Subjects were told to imagine they had $15 and that they could choose any fraction of that money to “send” to us. If we judged their work favorably, we would double what they sent and “send it back”—in the event of an unfavorable judgment, we would keep whatever they sent.\(^\text{13}\) The instructions and interface are shown in Appendix B, Figure 7. To save money, we told subjects that we would select one player at random and implement his or her choices. Although this game measures the worker’s trust in us (e.g., trust that we will follow our own protocol and judge their work fairly), it is confounded with the worker’s confidence, risk aversion and beliefs about our standards for the work. If these non-trust factors held constant across experimental groups, we could in principle detect changes in trust, though this assumption of constant effects is questionable. In any event, these factors should be unaffected by the treatment for the subjects in a cell we created in which subjects proceeded immediately to the games.

We measured cooperation in two ways: (a) willingness to give free advice and (b) willingness to take a short survey at a later date. For the first cooperation measure, subjects were asked to suggest ways to make the transcription task easier. The wording of the request made it clear that offering advice was vol-

\(^{13}\)The subject we randomly selected had chosen to send the full $15; we gave him or her $30.
untary and that no payment would be made for the offered advice.\textsuperscript{14} The instructions and interface are shown in Appendix B, Figure 10. For a second cooperation measure, we offered subjects the opportunity to take a follow-on survey for an additional 15 cents.\textsuperscript{15} The survey measure of cooperation captures a mix of self-interest and cooperation with the firm, and is somewhat analogous to a willingness to work overtime. The instructions and interface for the survey request are shown in Appendix B, Figure 8.

A.2 Consummate performance results

In addition to simply not accepting a wage cut or doing a poor job, workers who face wage cuts could withhold “consummate” performance. Consummate performance is difficult to measure—at least in some models, the difficulty of measuring (or at least contracting upon it) is definitional. These challenges aside, we attempted to measure consummate performance with the “games” phase of the experiment, which followed treatment. Each game was designed to give workers the chance to evince some changed attitude towards us that a real employer might find consequential. The games are all described in Section A. To aid in the comparison across groups, we created a cell in which subjects were not exposed to an offer to perform an additional paragraph transcription.

**ONLYGAMES:** Subjects were not offered an additional task. After the initial transcriptions, they went straight to the follow-on games.

We had hoped that our consummate measures would all be highly correlated with each other, suggesting that the different measures were all manifestations of the same positive or negative change towards us as an employer/trading partner. However, there is only weak evidence that this hope was borne out: Table 5 shows the correlation matrix for the four consummate performance measures, pooled across all experimental groups. We see that the patience measure had no strong relationship to any of the other measures and is even slightly negatively correlated with the advice measure. The advice and trust

\textsuperscript{14} The text of the actual question was, “Please help us make these tasks easier for other workers. Describe what steps workers can take to make doing these transcriptions easier.”

\textsuperscript{15} When you pay workers on MTurk, you can embed a short message about why you are paying them—we embedded the link to our survey in this message. Remarkably, not one person who agreed to do the survey actually did it after we sent them the link. We had intended to use completing the survey as another measure of cooperation and trustworthiness, but there was no variation in survey response. Given the universal rejection, we fear that some technical problem may have prevented completion.
measures are positive related, as are survey and advice measures—but survey and trust are not related to each other.

Table 5: Correlation matrix for consummate performance measures

<table>
<thead>
<tr>
<th></th>
<th>advice.z</th>
<th>trust.z</th>
<th>survey</th>
</tr>
</thead>
<tbody>
<tr>
<td>advice.z</td>
<td></td>
<td>0.17*</td>
<td></td>
</tr>
<tr>
<td>trust.z</td>
<td>0.17*</td>
<td></td>
<td>0.01</td>
</tr>
<tr>
<td>survey</td>
<td>0.15*</td>
<td>0.01</td>
<td></td>
</tr>
<tr>
<td>patient</td>
<td>-0.07</td>
<td>0.10</td>
<td>0.07</td>
</tr>
</tbody>
</table>

The first question about the any of the games is whether simply being asked to perform an additional paragraph transcription had any effect. For this, we can compare CONTROL, which received a 10 cent offer, to the ONLYGAMES group. Column (1), Table 6 reports an estimate of

\[ \text{SURVEY} = \beta_0 + \beta_1 \text{CONTROL} + \epsilon \]  

where the sample consists of subjects in CONTROL and the excluded group is ONLYGAMES. We can see that subjects that were presented an additional fourth paragraph were about 9 percentage points less likely to agree to a survey compared to those simply brought directly to the games. However, this estimate is not conventionally significant. Note that among ONLYGAMES, agreement to do the survey was 100%.

For the other games, we also find no evidence of a differential performance: in two-sided t-tests for both the advice z-score and the trust z-score, we fail to reject the null hypothesis of no difference in group means, with \( p = 0.29 \) and \( p = 0.174 \), respectively. Furthermore, the CONTROL coefficient is positive for the trust and advice measures.

Next we test whether any wage cut—regardless of how framed or justified—affect willingness to take the survey. Column (2) reports an estimate of

\[ \text{SURVEY} = \beta_0 + \beta_1 \text{ANYWAGECUT} + \epsilon, \]  

where ANYWAGECUT is an indicator that the subject was assigned to WAGECUT of any of the four wage cut groups. The comparison group is the CONTROL group, to which the 10 cent offer was made. The coefficient on ANYCUT is very close to zero, implying no detectable effect on willingness to take the
survey. As before, for the other games, we also find no evidence of a differential performance: in two-sided t-tests for both the advice z-score and the trust z-score, we fail to reject the null hypothesis of no difference in group means, with \( p = 0.26 \) and \( p = 0.67 \), respectively.

Table 6: The effects of wage cuts on consummate performance measures from the “games” phase of the experiment

<table>
<thead>
<tr>
<th>Dependent variable:</th>
<th>Agree to complete survey?, (SURVEY = 1)</th>
<th>Advice Z-Score</th>
<th>Trust Z-Score</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>(1)</td>
<td>(2)</td>
<td>(3)</td>
</tr>
<tr>
<td>CONTROL</td>
<td>−0.087</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.063)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Any wage cut offer?</td>
<td>−0.013</td>
<td>−0.101</td>
<td>−0.420</td>
</tr>
<tr>
<td></td>
<td>(0.067)</td>
<td>(0.092)</td>
<td>(0.312)</td>
</tr>
<tr>
<td>FAIRNESSPRIMEDWC</td>
<td>−0.202**</td>
<td>−0.202</td>
<td>−0.302</td>
</tr>
<tr>
<td></td>
<td>(0.094)</td>
<td>(0.094)</td>
<td>(0.317)</td>
</tr>
<tr>
<td>NEIGHBORWC</td>
<td>0.059</td>
<td>0.059</td>
<td>0.167</td>
</tr>
<tr>
<td></td>
<td>(0.092)</td>
<td>(0.092)</td>
<td>(0.312)</td>
</tr>
<tr>
<td>PRODUCTIVITYWC</td>
<td>0.002</td>
<td>0.020</td>
<td>−0.400</td>
</tr>
<tr>
<td></td>
<td>(0.091)</td>
<td>(0.091)</td>
<td>(0.309)</td>
</tr>
<tr>
<td>PROFITWC</td>
<td>−0.025</td>
<td>−0.025</td>
<td>−0.473</td>
</tr>
<tr>
<td></td>
<td>(0.093)</td>
<td>(0.093)</td>
<td>(0.314)</td>
</tr>
<tr>
<td>UNFRAMEDWC</td>
<td>1.000***</td>
<td>0.913***</td>
<td>0.941***</td>
</tr>
<tr>
<td></td>
<td>(0.046)</td>
<td>(0.062)</td>
<td>(0.071)</td>
</tr>
</tbody>
</table>

Comparison Group: ONLYGAMES, CONTROL, WAGECut, WAGECut, WAGECut

<table>
<thead>
<tr>
<th>Observations</th>
<th>ONLYGAMES</th>
<th>CONTROL</th>
<th>WAGECut</th>
<th>WAGECut</th>
<th>WAGECut</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>44</td>
<td>163</td>
<td>140</td>
<td>140</td>
<td>140</td>
</tr>
<tr>
<td>( R^2 )</td>
<td>0.043</td>
<td>0.0002</td>
<td>0.085</td>
<td>0.057</td>
<td>0.020</td>
</tr>
</tbody>
</table>

Notes: This table reports OLS regressions where the dependent variable is an indicator for whether the subject agreed to complete a short survey in the future, in Columns (1) through (3) and a z-score measure of the workers offered advice, in Column (4) and their trust in us as employers/trading partners, in Column (5). In Column (1), the comparison group are those subjects assigned to ONLYGAMES, which were not offered the chance to do a fourth transcription. They are compared against CONTROL, which was offered 10 cents. In Column (2), all cells where a wage cut was proposed (i.e., 3 cents for a fourth paragraph) are compared to the CONTROL to see if collectively the wage cut lowered willingness to take the survey. Finally, in Column (3), each of the justified or primed cells in which 3 cents was offered is compared against WAGECut, the unexplained wage cut. Robust standard errors are reported. Significance indicators: \( p \leq 0.10 : * \), \( p \leq 0.05 : ** \) and \( p \leq 0.01 : *** \).

Despite no overall effect of a wage cut, we know that the different justifications strongly affected willingness to accept the wage cut. To test whether these justifications affected consummate performance, Column (3) reports a regression where ANYWAGECut is decomposed into each of the distinct justifications, with WAGECut as the comparison group. In this regression, we see that the coefficient on PROFITWC is negative and large, with a nearly 20 percentage point reduction in agreement to complete the survey. The effect is conventionally significant. However, given the multitude of treatment cells, the
finding of one significant effect out of five could very likely be due to chance. On the other hand, finding the strongest negative effects in the cell with the most unreasonable justification could indicate an actual behavioral response. And yet there is additional evidence for the argument that the seeming effect of the profits treatment is spurious: in Columns (4) and (5), we change the outcomes to the advice and trust z-scores and find no substantial differences across cells. For both outcomes, we would fail to reject the null hypothesis of zero coefficients on all group indicators ($p = 0.16$ for the advice measure and $p = 0.74$ for the trust measure). However, for both the trust and advice measures, we are under-powered to detect anything but very large effects, with the standard errors being more than $3/10$ths of a standard deviation in the outcome variable.

B Experimental materials

The actual surveys used for each of the experimental groups are hosted at the following URLs:

- **CONTROL**: http://www.surveymonkey.com/s/R23DT7Y
- **WAGECut**: http://www.surveymonkey.com/s/R23FZHW
- **ONLYGames**: http://www.surveymonkey.com/s/R2G53NM
- **PRODUCTIVITYWC**: http://www.surveymonkey.com/s/R2BXXM7
- **ProfitWC**: http://www.surveymonkey.com/s/R2H6QQ2
- **NeighborWC**: http://www.surveymonkey.com/s/R2HMWZR
- **UNFRAMEDWC**: http://www.surveymonkey.com/s/R2699DW
- **FAIRNESSPrimedWC**: http://www.surveymonkey.com/s/R2Z2SH9
Notes: This was the first page that all participants—regardless of experimental group—first landed on.
Figure 5: Transcription environment for all paragraphs

Notes: All subjects transcribing paragraphs used this interface for both the initial three paragraphs and the additional paragraph. Note that the task value differed depending upon the treatment assignment.

Figure 6: G8 Fairness priming question

Notes: Subjects in G8 were asked what was a fair wage for a paragraph, using this dialog.
Figure 7: “Trust game” dialog

We are interested in learning how confident people are in the quality of their work.

We will select one of you at random to be given $15. If you are the one to receive $15, you can choose some amount between $0 and $15 to send back to us. Suppose you send $x dollars. If your work is judged to be very good, we’ll double x (i.e., pay you 2x) and you will still keep the 15-x you did not send. If your work is judged to only be mediocre, you will lose the x, but still keep whatever you did not send.

Example:
1. Sally decides to send $10. Her work is judged to be only average - she loses her $10 but keeps $5.
2. Dave sends less than Sally. He sends nothing. His work is judged to be only average, but he loses nothing and keeps the whole $15.

We will select one worker at random, evaluate their work and then implement their choices. If you are the selected person, you will be paid a bonus ranging from $0 to $30, depending upon your choices and our evaluation. Assume we have given you $15.

How much would you send?

Notes: This screenshot shows the instructions and choices for subjects playing the contextualized “trust game.”

Figure 8: Follow-on survey dialog

1. Would you be willing to fill out a survey that will be sent as a link with your bonus? We would pay you an additional 15 cents.

   - yes
   - no

Notes: This screenshot shows the interface in which we presented subjects with the opportunity to take a follow-on survey at a later date.

Figure 9: “Patience” dialog

1. Would you be willing to be paid in two weeks for an additional bonus of 30 cents so we can wait until all the transcriptions are done?

   - yes
   - no

Notes: This screenshot shows the interface in which we presented subjects with the opportunity to delay getting paid in exchange for a larger payment at a later date.
Figure 10: “Advice” dialog

Notes: This screenshot shows the interface in which we asked subjects to provide advice to other workers.

Figure 11: Demographics dialog

Notes: This screenshot shows the interface in which we asked subjects to provide some demographic information. It also shows the completion code link that subjects used to generate a code which could link their responses to their MTurk identities.
Figure 12: MTurk tasks that Amazon supports via pre-made templates and vetted workers

Notes: This is a screenshot of the Amazon Mechanical Turk interface for starting a new project from a template. The column of text on the left side of the image shows the common use-cases for MTurk. The “Transcription from an Image” task is selected and shown, which is the kind of task we used for our experiment.