A Normative Methodology for Modeling Consumer Response to Innovation

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Consumer response determines the success or failure of new products and services. This paper proposes a methodology that integrates knowledge in the fields of psychometrics, utility theory, and stochastic choice theory to improve the design of new products and services. The methodology consists of a consumer response and a managerial design process. The design process is one of idea generation, evaluation, and refinement, while the consumer response is based on consumer measurement, models of the individual choice process, and aggregation of predictions of individual choices. The individual response model processes the consumer measures by first reducing them to an underlying set of perceptual dimensions. Then the measures of perception are combined to produce a scalar goodness measure for each choice alternative through a process called "compaction." Next, homogeneous segments are defined based on similar preferences. The goodness measures for each consumer or segment are linked to probability of choice for the new products and services and for competing products and services. In each step theoretical, empirical, and statistical issues are identified. Various techniques are introduced and described for each phase. Selected techniques are demonstrated based on the survey data collected at MIT to support the design of a health maintenance organization (HMO) and in the consumer market to evaluate a new deodorant.

Almost all organizations face the problem of how to develop and introduce successful new products or services. This problem receives high strategic importance since such innovation is linked to increased effectiveness and productivity. In the private firm successful new products result in sales and profit growth. For example, approximately 50% of the growth in sales over a five year period in many industries was accounted for by new products (Booz, Allen, and Hamilton [1]). In services like transportation, additional passengers, efficiency, and revenue may be obtained by new services such as computer-controlled mini-buses. Innovation in
the design of service packages in the fields of insurance and finance can improve the competitive positions of companies and ensure a stable base for corporate growth. In the field of health, the health maintenance organization (HMO) provides an example of an innovative new service. A successful HMO could have lower costs and higher quality of care along with high enrollment and re-enrollment. Although the measures of effectiveness vary across public and private industries, new products and services are critical to vital functioning and achievement of goals.

While new products and services are crucial to organizational growth and effectiveness, they also represent a high risk to the organization. Many new products fail. Approximately 30% of the new products introduced by firms in the market fail and 80% of the resources for new product development are allocated to products that are not a success in the market [1]. In public organizations many failures have been recorded in public programs such as low-cost housing, mass transit services, and preventive health services. Many of these failures reflect a lack of acceptance by consumers. The products did not sell enough or the public services were not used by the clientele.

The critical role of the consumer in the management of innovation is being more clearly recognized. Private firms that sell directly to the mass market have long recognized that understanding consumer needs is the key to successful innovation. New evidence indicates that even in high technology areas 60-80% of successful technological innovations are generated by consideration of consumer needs (Utterback [63]). Since it is clear that most successful innovation in private firms is due to understanding consumer needs and responses, it is probably reasonable to posit that this same effort could increase the rate of success of innovations in public organizations as well.

Private firms allocate substantial effort and resources to developing new products through research and development and marketing departments. Figure 1 depicts a process for development of innovative products and services. The first step is design. In this step consumer studies are integrated with technology and creative efforts to generate new ideas. These ideas are then evaluated and refined on the basis of consumer reactions, production issues, and financial considerations. After an idea has been established as a viable and significant innovation, it is tested in a pilot program or test market. If the test is successful, the product can be introduced.

This paper will address the problems of the design stage of new product development. Emphasis will be on integrating consumer response into the design activities of idea generation, evaluation, and refinement. This integration will be done through a behavioral-process model of individual response to innovation. After summarizing the most relevant existing work
in the fields of psychometrics, utility theory, and stochastic choice theory, we will define the macro model structure. Next the measurement, estimation, and micro structural issues will be discussed. We will provide specific examples based on: (1) the problem of designing a new prepaid, comprehensive health service plan (HMO) and (2) consumer evaluation of a new deodorant. We conclude with a description of future research needs.

![Design Process Diagram]

**Figure 1.** Process for development of new products and services.

1. **SOME EXISTING WORK**

**Psychometrics**

Psychometricians are concerned with the problem of how individuals perceive stimuli. Using measurements of perceived similarity among stimuli and measurements of attributes for new and existing stimuli, perceptual maps can be developed by multi-dimensional scaling procedures (Kruskal [37], Young and Torgerson [68]). These perceptual maps identify the important dimensions that consumers use to distinguish between stimuli and indicate the position of each stimulus relative to these dimensions. In marketing the stimuli are products and the map defines market structure. Opportunities for new products are identified by examining the gaps in the market structure (Stefflre [56], Green and Carmone [14]).

Preference judgments can be integrated with the perceptual data to indicate high opportunity areas. PREFMAP is a popular method for ac-
completing this task (Carroll and Chang [3], Carroll [2]). PREFMAP uses regression to derive an "ideal" point and relative importances of the dimensions from stated preferences of the consumers regarding the existing stimuli. Srinivasan and Shocker [55] use linear programming in an alternative fitting procedure for estimating importances. Another approach is through conjoint analysis (Tversky [59]), which draws on an axiomatic and statistical base to produce relative importances by requiring consumers to rank order preferences for factorially generated combinations of product attributes (Green and Wind [17], Johnson [25]).

While PREFMAP and conjoint analysis use statistical procedures to impute the importances, other psychologists use "expectancy value" models that use direct consumer judgments to estimate importances. Extensive work has been done on such models based on psychological theories of attitude formation (Fishbein [9], Rosenberg [49]) and their extensions (Ryan and Bonfield [50]). Most of the models are conceptually similar in that they define an attitude toward an object as a linear additive function of an individual's reactions to an object on an attribute scale multiplied by a measure of the effect of that attribute in the overall attitude formation. A common formulation is the linear combination of the "importance" of each attribute multiplied by the individual's belief as to the extent to which the attribute is offered by a specific alternative (Wilkie and Pesse- mier [66]).

Utility Theory

While the psychometricians apply a methodology based on multidimensional scaling and statistical preference analysis, utility theorists approach a similar problem from a substantially different point of view. Prescriptive utility theory is oriented toward helping managers make policy decisions under uncertainty and derives its strength from a rigorous set of axioms (von Neumann and Morgenstern [84]) and theorems that specify unique functional forms, e.g., additive, multiplicative, and quasi-additive (Raiffa and Schlaifer [47], Keeney [28, 29, 31], Richard [48], Farquhar [7, 8], and Fishburn [10]). The coefficients of these functions reflect the relative importances of the relevant performance measures, their interdependencies, and the risk averseness of the decision maker. The theorems also indicate techniques to assess and test the preference parameters by asking individuals to state when they are "indifferent" between two alternative stimuli. The dependent value of the utility function is a single cardinal measure of goodness of an alternative. Since the theory is used to guide the decision rather than describe it, the decision maker chooses the alternative with the highest expected utility value.

Most empirical applications have been based on directly assessing the utility function of one or a small number of decision makers based on a set
of quantifiable attributes of alternatives (Keeney [30]). This is in contrast to the psychometrician's approach, which is based on interviewing many consumers on perceived attributes that must be individually scaled.

**Stochastic Choice Theory**

Recognizing that there will always be uncertainty in any prediction of choice behavior, economists, transportation demand theorists, and mathematical psychologists concentrate on axioms to determine selection probabilities from observable "scale" values (Luce [39], McFadden [43]). Economists and demand theorists parameterize scale functions and statistically estimate the parameters from observations on actual choice among existing alternatives. Popular models for this are the multinominal logit and other "random utility" models (McFadden [42]).

Mathematical sociologists model the stochastic choice process directly through diffusion, learning, Bernoulli and semi-Markov models (Massy, Montgomery, and Morrison [41], Coleman [5]). These models describe the dynamics of choice probabilities over time but do not link attributes of products or consumer preferences to choice.

**Discussion**

Although a good deal of work is being done, it is clear that the work is very diverse. Each discipline reflects different measurement approaches, analytic techniques, and foci. Psychometricians are concerned with perceived attributes and recovery of importances from stated consumer perceptions and preferences. Utility theorists are concerned with theoretical soundness of functional forms through axiomatic consistency and with direct assessment of relative importances, attribute interdependence, and the risk characteristics for the purpose of aiding decision making. Choice theorists axiomatically model linkages to probability of choice but do not consider linkages between consumer perception and managerial prediction of attributes or axiomatic specification of their functional utility forms.

The approaches also differ in how they treat the issues of aggregation. Most psychometricians develop average representations of perception and preference but explicitly check that they are homogeneous with respect to perception and preference (Carroll and Chang [4], Tucker and Messick [58]). Utility theorists and conjoint analysts work completely idiosyncratically. Demand choice theorists can directly model individual response, but their statistical techniques force judgmental specification of aggregate segments before parameter estimation.

Although the approaches are incomplete and diverse, they are complementary, with each being primarily directed at a different phase in the consumer-choice process. We visualize a process of perception, preference,
and choice that integrates the approaches to form a complete consumer-response model. Some initial work has been done to integrate these disciplines, but only at an aggregate level (Urban [62], Pessemier [44]).

2. MACRO DESCRIPTION OF THE METHODOLOGY

This section proposes a methodology that draws on the existing work in psychometrics, utility theory, and stochastic modeling. It attempts to be comprehensive by integrating existing approaches into a cohesive but modular process that offers a variety of techniques of varying complexity and data requirements. It uses structures that reflect the acceptance phenomena at a level consistent with what is known about behavior. Efforts are made to make assumptions explicit, isolate weaknesses in existing techniques, indicate where improvements need to be made, and prevent models from being used in applications that violate their assumptions. Attention is focused on models that can predict response to changes in design and that can be extended to design changes or to new alternatives that are outside of existing consumer experience. For example, the models make predictions of consumer response to a new HMO, even though none currently exists in the community.

Since the ultimate value of the methodology will be in better design of products and services, creativity is recognized as a critical element. The methodology elicits and focuses creativity by identifying characteristics relevant to the choice process and by explicitly measuring relative importance of these characteristics. Although some steps are technically complex, the underlying choice-process structure is understandable to nontechnical as well as technical members of a product team. We attempt to make the outputs of each step clear and understandable so that the design team can visualize the choice process and can create and refine new products or services.

Design Process

The methodology is graphically illustrated in Figure 2 as a managerial design process and a parallel consumer-response process. The analytics and the focus of this paper are on the consumer-response process. First measures of perception and preference with respect to the relevant choice alternatives are observed for a sample of consumers. These measures are used to estimate the parameters of a model of the individual choice process. Finally, an estimate of group response is obtained by aggregating individual acceptance measures (probabilities). The measures of group response are then input to the evaluation model in the managerial design process, which includes consideration of investment, operating costs, risk, and externalities as well as consumer acceptance.

It is rare that a new product or service will be implemented on the basis
of a single cycle through the methodology. Instead a screening process will result that identifies the most promising alternatives for further considera-

Figure 2. Components of consumer response methodology.

tion. These alternatives are refined on the basis of detailed diagnostic information generated by the individual choice models (see Arrow B in Figure 2). The refined design can be analytically tested in the individual choice models and the simulated results can be used iteratively to lead towards a "best" design (see arrows A and B in Figure 2). This "best"
design identified by the iterative process is then tested by taking new consumer measures and cycling through the entire methodology (arrows C and D). In the early design phases attention is upon design specification and improvement (arrows A and B), while in the later phases of design, attention is focused on evaluation and refinement (arrows C and D).

In the early phases of a new product or service design, measures can be obtained only on concept descriptions. As the product or service design evolves, more comprehensive descriptions become possible until in the final phases the product concept is a real choice alternative executed to the stage of advertising copy, package design, price, promotion, and distribution strategy. In this paper we present an early design problem for a new health service. Since additional models and measurement become necessary in later design phases, we illustrate these with an example from the evaluation of a new, frequently purchased consumer product.

We begin our discussion with a brief overview of the consumer response model. The concepts, models, measures, and notation will later be discussed in detail and illustrated with empirical examples.

**Consumer Response Components (Overview)**

The basic input for the consumer response model is generated by surveys of the potential users of the new product or service alternatives. This set of information is denoted by Ω. The individual choice process consists of the analytic phases of (1) perception, (2) “compaction,” (3) segmentation, and (4) probability of choice (see Figure 2). The consumer response process is modular because this structure allows more effective integration of the disciplines of psychometrics, utility theory, and stochastic choice theory, which are each closely associated with one module. In the perception phase the attitude evaluations of choice alternatives in measurement set (Ω) are reduced to a smaller set of underlying perceptual performance dimensions. Emphasis is on the designation of the number of underlying dimensions and their names. This reduced set of perceptions is represented by X and is made up of each consumer’s (i) perception of each performance dimension (m) for each choice alternative (j).

Next these multiple measures of perception (X) are combined to yield one measure of the goodness for each alternative. We call this operation “compaction” because the several perceptual measures of an alternative are “compacted” into one measure of evaluation. We defined a new descriptive word since the procedures of utility theory, PREFMAP, expectancy value models, and conjoint analysis are all directed at this one task. In the compaction phase, the vector of individual performance scale values for each alternative for each dimension ((xij)=(xij1, ..., xijM)) and a vector of individual parameters (ζi) are variables in a real-valued function c(xij, ζi), which compacts them into a scalar measure of good-
ness \( c_{ij} \). A separate goodness value is determined for each individual \((i)\) in the sample and for each of his choice alternatives \((a_j)\).

On the basis of the preference parameters \((\lambda_i)\) of the compaction functions, homogeneous groups of consumers are abstracted for designation as segments \((s)\) of the target population. Within each segment \((s)\) distributions of the performance dimensions \((X_i)\), the preference parameters \((\Lambda_s)\), and the functional form of the compaction function \(c_s(x_{ij}, \lambda_i)\) are determined. This segmentation is specific in its criteria of homogeneity of the preference parameters \((\lambda_i)\) within the segment. For example, consumers who have similar importances for each perceptual dimension are grouped together. Specific procedures and explicit tests for segmentation are discussed later and compared with existing segmentation methods.

Empirically, we observe that consumers do not always choose the alternative with the highest scalar measure of goodness. In this methodology the scalar values are considered as independent variables in a probability of choice model that links an individual's vector of goodness measures \((c_{a1}, \ldots, c_{a\ell})\) to his choice probabilities \((p_{ij})\). Each individual choice probability \((p_{ij})\) for each alternative \((a_j)\) is derived by a function \((p_{ij}(a_j|c_{a1}, \ldots, c_{a\ell}))\). The subscript \(s\) indicates that the functions, but not the probabilities, are the same for all individuals in segment \(s\).

The final step aggregates the individual choice probabilities to obtain group response measured by the mean \((\bar{m}_s)\) and variance \((\bar{m}_s)\) of share of choice, or in some cases the mean \((\bar{N}_s)\) and variance \((\bar{N}_s)\) of the total number of people choosing each alternative. If it is managerially useful, aggregation can be done separately within segments. By explicit modeling of the various stages of consumer response, useful diagnostics can be identified for the refinement model. These diagnostics help the managers understand the market structure, segmentation, relative importances and interdependence of product attributes, and the risk characteristics of the choice process. The model itself processes individual information, but aggregate results from perception, segmentation, and compaction are important in eliciting creativity for the design of innovative alternatives. For example, in the perception phase, average perceptual maps of product positioning, and in compaction, average values of the importance weights guide the manager in effective design of new products or services. These are particularly useful in the early design phase, where the product must be positioned in the space of perceived product attributes. In the late design phase such positioning is of less concern, and attention focuses on probability of choice and accurate forecasting and evaluation.

3. MICRO ASPECTS OF THE METHODOLOGY

This section investigates each module of the consumer-response process and gives examples of the measurement and estimation procedures. The
Figure 3. Guide to empirical presentation of consumer response process.
major source of the examples is a study of health care innovation, but because
the health care study (HMO study) concentrates on the early design process, some of the modules are further illustrated with measures from the study of a new deodorant. We give Figure 3 now to enable the reader to view each example in its relationship to the total design process. Various techniques will be discussed throughout the paper, but all the in-depth empirical examples are shown in Figure 3. In selecting empirical examples, we emphasized new methods and techniques.

Consumer Measurement

A good model is dependent upon high-quality input. A model is accurate only if the measurements it requires are valid and reliable. This section discusses the measurement issues of the methodology. The HMO measurement issues will be presented in some detail and then the deodorant data will be briefly discussed.

Early Design—HMO: Two samples were drawn from the target population for a new health maintenance organization at MIT. The first sample included faculty, students, and staff and was used for the perception, statistical compaction, and segmentation phases of the study. The second sample was drawn from the student subset of the target population and was used to test the feasibility of applying von Neumann-Morgenstern utility theory for directly measuring compaction functions.

Consumer perceptions, preferences, and choice behavior must be observed with respect to the product or service alternatives that are “evoked.” For example, there may be a large number of products available, but each consumer only “evokes” a few of these. In a study of seven consumer products, consumers had an average of only three “evoked” brands. “Evoked” was defined as brands last used, ever used, on hand, or would not consider using [62]. In services, especially new ones, the number of alternatives is often so small that one must force evoking by the use of concept statements in order to have sufficient perceptual inputs. For example, in the study of HMO design at MIT the only real option available was existing private care. Thus the evoked set was expanded to four options by specifying three new options in concept form: an MIT HMO, the Harvard Community Health Plan, and a hypothetical Massachusetts Health Foundation. See Table I for an example of the MIT HMO option. One thousand surveys were mailed to a random sample of the MIT community and 447 faculty, students, and staff completed and returned the questionnaire.

First a set of important product attributes are identified and consumers evaluate each evoked choice alternative with respect to each attribute. These attributes and their descriptions are generated from in-depth in-
terviews with individual or groups of consumers or by Kelly's [33] triad procedure, in which consumers describe how the two most similar of three stimuli are alike and how the two most dissimilar are different. Consumers then rate the attributes on bipolar or agree/disagree scales. In the HMO

| TABLE I |
| HMO Concept Description |

DESCRIPTION OF THE M.I.T. HEALTH PLAN

M.I.T. announces a new health care plan for YOU AND YOUR FAMILY. By joining the M.I.T. HEALTH PLAN you can get comprehensive health care at a low, fixed monthly charge. Virtually all your medical needs will be met. You will not have to face unexpected doctor or hospital bills and you will not have to worry about finding a good doctor for you or your family.

The cost of joining the M.I.T. HEALTH PLAN is only a little more than regular Blue Cross/Blue Shield health insurance, but you get more services and comprehensive care. There are no charges for doctor visits, nursing and laboratory services, or hospital services. Women in the plan pay nothing extra for prenatal, delivery, or maternity care. The services are comprehensive and include mental health care and emergency services.

The costs are kept low by the utilization of preventive care to keep you well. The plan succeeds by keeping you and your family well and out of the hospital. In addition, the use of trained paramedicals and technology helps reduce costs while maintaining the quality of care.

You choose your own personal doctor (specialist in internal medicine for yourself and a pediatrician for your children) from our staff of physicians. Your doctor supervises your total health care at the health center and in the hospital. He will be sure you get the highest quality of care. When you are a member of the M.I.T. HEALTH PLAN you can be sure of getting health care around the clock from the staff of physicians, nurses, social workers and allied health personnel.

The M.I.T. HEALTH PLAN delivers its services from the Homberg Memorial Building on the M.I.T. campus. Parking is available during patient visits. Hospital services are provided by the Mount Auburn and Cambridge City Hospitals. Maternity and gynecology care are provided through the resources of the Boston Hospital for Women. For emergencies outside the Boston area, local hospitals can be used.

You can become a member of the plan by paying $1.50 per month more than your Blue Cross/Blue Shield coverage if you are single and $4.00 more per month if you are married. If you are a single student and do not have hospital insurance, the cost is $8.25 month more than the student health fee you are currently paying; if you are a married student, the cost is $20.00/month more than the student health fee. These fees cover all of your medical costs except: the first $50 and 20% of the balance of prescription charges and the excess of $10 per visit for psychotherapy (over $5 per visit for group therapy). The plan does not include eye glasses, hearing aids, cosmetic surgery, custodial treatment, or dental care done outside a hospital. If you join the plan, you must remain a member for one year.

The M.I.T. HEALTH PLAN is designed to make comprehensive, high quality health care available to you and your family at a low cost.
study 16 statements that consumers had earlier defined as relevant to their health care were evaluated by 5-point agree/disagree scales (see Table II).

After perceptions have been determined, rank order or constant sum preference measures are obtained for use in statistical compaction techniques. Where possible, constant sum paired comparisons are preferred since they yield interval scales (Torgerson [57]). The initial HMO survey was done by mail and hence the easier-to-answer rank order preference measures were collected. Consumers were then asked to choose among alternatives. In the HMO study consumers were given choices (0, 1) between existing care and the HMO. If they chose MIT, a 5-point intent scale was administered. Then other alternatives were added to the choice set and new intent measures taken.

While the procedures for measurement of perception and preference are comparatively well developed, this is not true in utility theory. There are no reports of measurement of consumer utility functions, partly because past uses of utility assessment were oriented as prescriptive decision applications with one or few decision makers (Keeney and Raiffa [32]). When consumers are considered, two assessment issues must be addressed:

1. How can perceptual phenomena be integrated in the measurement?
2. Is it possible to have consumers understand the required lottery and trade-off questions and give meaningful answers?

When directly assessed utility models are to be supported (see compaction section of this paper for a more detailed description of the procedure), data are required to measure risk averseness, importances, and interactions relative to various "performance measures." Prescriptive utility theory requires these to be quantifiable, instrumental variables such as cost or waiting time, rather than perceptual measures like quality or "personallness" of health care. The former are easier for the manager to relate to, but the latter better reflect the consumer choice process. In this methodology we propose that the psychological dimensions obtained by reducing the perceptions of choice alternatives be used as performance measures.

In the HMO case a special study was conducted to assess the utility functions of 80 students. First 16 attributes were factor-analyzed to obtain four perceptual measures (see the reduction section of this paper for a more detailed description of the procedures). The dimensions were "quality," "personallness," "convenience," and "value" and the 80 students additionally rated the four health alternatives directly on 7-point scales for these four reduced perceptual dimensions. Utility functions were then assessed relative to the performance measures defined by the perceptions of the 7-point scales. After utility assessment the overall 7-point scale values can be correlated to the factor scores and therefore to the original
<table>
<thead>
<tr>
<th></th>
<th>Strongly agree</th>
<th>Agree</th>
<th>Neither agree nor disagree</th>
<th>Disagree</th>
<th>Strongly disagree</th>
</tr>
</thead>
<tbody>
<tr>
<td>1.</td>
<td>a</td>
<td>b</td>
<td>c</td>
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<tr>
<td>2.</td>
<td>a</td>
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<td>3.</td>
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<td>5.</td>
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<td>10.</td>
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<td>12.</td>
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<td>13.</td>
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<td>14.</td>
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<td>15.</td>
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<td>16.</td>
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(Values assigned = $-2 -1 0 +1 +2$)
perceptual ratings. By defining performance measures through perceptual scaling methods, utility theory can be meaningfully linked to psychometrics.

The use of a reduced number of perceptual dimensions as performance measures also makes consumer utility measurement more feasible. Since the number of dimensions is small (usually four or less), measures of risk aversion and importance need be collected only on this smaller set of performance measures.

Risk aversion is measured by having the consumer consider a lottery on the performance measures. Although it was anticipated that this would

**Figure 4.** Schematic of risk aversion question.

be a difficult task for consumers, it was found in the pilot study that the 80 students related well to a carefully designed questionnaire if simple lottery questions were included to educate them to the task required and the meaning of a probability. In fact, the only difficulty was with students already well schooled in probability theory who tried to give expected value answers rather than their true feelings. The procedure is schematically represented in Figure 4. The respondent sets the area of a probability wheel so that he would be indifferent between the certain health plan outcomes and uncertain health plans as represented by lottery outcomes. Most respondents were comfortable with this task and all completed the interview. Utility assessment requires two lotteries for each performance measure or a total of eight lottery judgments in the HMO case. One lottery, where one performance measures varies while all others are
held fixed, is enough to determine the risk characteristics of that performance measure if constant risk aversion and utility independence are assumed (Raiffa [46], Keeney [31]). The second lottery is needed to verify behavioral assumptions inherent in the form of the utility function.

Relative importance weights are determined by asking consumers to trade-off one performance measure, say convenience, with another, say value, while holding all others fixed (see Table III). Another trade-off question then varies the fixed values to verify a behavioral assumption known as preferential independence, which, together with utility inde-

**TABLE III**

**Schematic of Trade-Off Question**

Instruction to Consumers:

Now consider the two plans below and choose the level of the quality factor in such a way that you are indifferent between the two plans. (Consumer is challenged and the question iterated until a true indifference is determined).

<table>
<thead>
<tr>
<th>Plan A</th>
<th>Plan B</th>
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<tbody>
<tr>
<td>Value = 5 (good)</td>
<td>Value = 5 (good)</td>
</tr>
<tr>
<td>Personalness = 5 (good)</td>
<td>Personalness = 5 (good)</td>
</tr>
<tr>
<td>Convenience = 6 (very good)</td>
<td>Convenience = 2 (very poor)</td>
</tr>
<tr>
<td>Quality = 2 (very poor)</td>
<td>Quality =</td>
</tr>
</tbody>
</table>

...

pendence, specifies uniquely the form of the utility function [31]. Thus two trade-off questions are required for all but one of the performance measures. This makes a total of six trade-off questions and eight unidimensional lottery questions in the HMO case. Finally, one complex lottery involving simultaneous changes in all the performance measures is used to measure interactions [32].

The results from the HMO study indicate perceptual dimensions can function adequately as performance measures and it is feasible, at least in the case of students, to measure consumer risk aversion and importances. Future research will assess the practicality of measuring utility parameters in a general population sample.

In both the campus-wide and student utility sample, demographics and other consumer descriptors were collected in order to adequately project from a sample population. For example, patterns of health care utilization
and satisfaction were measured in addition to demographics such as age, sex, family size, and health status.

**Late Design—Deodorant:** As the product or service design is finalized, the accurate forecasting of demand becomes more important and measurements are changed to simulate more closely actual environments. In the deodorant study consumers were intercepted in a shopping mall, exposed to TV advertisements for the new and old products, given the opportunity to buy the product from a retail shelf, and took the product home for use. The reader is referred to Silk and Urban [53] for a detailed discussion of the measurement design and execution.

Since a personal interview was used, constant sum preference measures [57] were collected before exposure to the new product and after home use. These before and after preference measures became the input to the probability of choice model that will be discussed later.

**Perception**

In the perception phase of the consumer response model, the attitude data collected in the measurement phase are reduced to a smaller underlying set of psychological dimensions. For completeness it is necessary to measure consumers on a large number of possible perceptual dimensions. But it is often difficult for managers and analysts to gain insight from comparison of perceptions relative to a large number of scales. For example, Figure 5 shows an average rating of consumers for the four HMO alternatives on the 16 scales shown in Table II. What the manager requires is a simpler representation that can be easily visualized and internalized for further processing. Thus in the reduction module of the consumer response model, the perceptual data are reduced to a smaller underlying set of psychological dimensions. These dimensions capture the essence of the perceptual process in a form that is readily understandable and more appropriate for use in design.

Several multidimensional approaches were available. If similarity judgments as well as ratings are collected, nonmetric techniques can be used to place the stimuli in perceptual space [14]. In many studies the evoked set is too small \( n \leq 8 \) or too varied across individuals to use nonmetric techniques to achieve statistical significance (Klahr [34]). To overcome this, the ratings can be directly reduced by using factor analysis on a data matrix in which each row reflects an individual’s rating of a stimuli. In this manner, even if each individual evokes only a small number of alternatives (e.g., \( n \leq 4 \)), the number of observations is large and equal to the number of individuals times the average number of choices evoked.

In all reduction methods care must be taken to test the results. In this methodology the sufficiency of the reduction is tested by correlations to preference and choice at later stages in the methodology. If the number of
1. I would be able to get medical service and advice easily any time of the day and night.

2. I would have to wait a long time to get service.

3. I could trust that I am getting really good medical care.

4. The health services would be inconveniently located and would be difficult to get to.

5. I would be paying too much for my required medical services.

6. I would get a friendly, warm and personal approach to my medical problem.

7. The plan would help me prevent medical problems before they occurred.

8. I could easily find a good doctor.

9. The service would use modern, up-to-date treatment methods.

10. No one has access to my medical record except medical personnel.

11. There would not be a high continuing interest in my health care.

12. The services would use the best possible hospitals.

13. Too much work would be done by nurses and assistants rather than doctors.

14. It would be an organized and complete medical service for me and my family.

15. There would be much red tape and bureaucratic hassle.

16. Highly competent doctors and specialists would be available to serve me.

Figure 5. Average ratings.

dimensions and their interpretation is inappropriate, the compaction phase preference prediction will be poor and provide a warning to the analyst. In the HMO study, principal component factor analysis was used to reduce the ratings on the sixteen scales for the four plans across 234 individuals
to four underlying dimensions. These four factors explained 55% of the total variance. Table IV presents the factor loadings (correlations) of the raw scales to the new underlying dimensions. By examining the high loadings on each dimension they were labeled judgmentally: (1) quality, (2) personalness, (3) value, and (4) convenience. Quality correlated to trust, preventive care, availability of good doctors, and hospitals. Personalness reflected a friendly atmosphere with privacy and no bureaucratic hassle. Value was not just price, but rather paying the right amount for the services. Convenience reflected location, waiting time, and hours of operation. A common factor analysis yielded similar interpretations. Factor scores were obtained that described the location of each plan on each dimension for each individual.

**Compaction**

The first analytic module in the methodology tells us how consumers perceive the alternative products or services, but it does not tell us what trade-offs consumers make in their decision to buy a product or select a service. To guide managerial decisions, the design team needs to know how...
consumers combine their perceptions on each performance dimension to form an evaluation of a product or service. For example, in an HMO should one increase the quality of the health service and charge a premium price or should the price be minimized subject to an adequate level of quality? Compaction generates this understanding by explicitly identifying the importance of each performance measure and its interaction in the consumers’ evaluation of an alternative. Formally this means that the performance measures identified in reduction are now compacted to form for each individual a scalar measure of goodness for each alternative. Preference and ultimately probability of choice result from comparison of these goodness measures.

Specifically, in compaction one determines a function \( c_i(x_{ij}, \lambda_i) \) that maps the vector of performance measures, \( x_{ij} \), into a scalar measure of goodness (a real number). The performance measures are the result of the reduction step in the methodology and the choice parameters result from the measured preferences or trade-offs and lotteries. For a given alternative, e.g., health plan \( a_j \), this scalar measure of goodness, \( c_{ij} \), has the property that with all other alternatives held fixed, any set of performance measures yielding the same value, \( c_{ij} \), must also yield the same probability of choice for alternative \( a_j \). In other words, compaction compresses the performance measures for an alternative into a one-dimensional measure, and knowing the value of this measure for each and every alternative is then sufficient to predict choice.

There are a number of techniques to estimate the parameters of a compaction function, but there are four basic categories: (1) direct consumer statement of importances, (2) statistical estimation of choice parameters \( (\lambda_i) \), (3) conjoint analysis, and (4) von Neumann-Morgenstern utility theory.

The most elementary technique is to ask consumers to scale the importance of each measure and form a compaction function as a linear additive function of an individual's reactions to an object on an attribute scale (e.g., rating on scale) multiplied by a measure of the effect (e.g., importance) of that attribute in the overall attitude formation. These models received considerable attention in marketing from psychologists (see [66] for a summary). Empirically, these models have been tested by correlation of the predicted attitude value with preference or choice. Empirical results have been mixed. Ryan and Bonfield [50] report correlations as high as 0.7 to 0.8 for an extended Fishbein model, while Sheth and Talarzyk [51] report correlations in the range of 0.1 to 0.4. Wilkie and Pessemier [66] in their review of 42 studies identify 19 with favorable results, 14 questioning the model, and 9 not applicable.

Another approach to compaction is to statistically estimate importances by a regression of observed preference against the perceptual attribute
measures. PREFMAP [2, 14, 16] is based on a regression of individual preference parameters for the consumers' evoked set of alternatives. In cases where the evoked set is small \((n \leq 8)\), PREFMAP has very few degrees of freedom to estimate individual parameters. PREFMAP also estimates an "average ideal point," but this is based on the average preference across individuals and is subject to the same degrees of freedom limitations. Furthermore, it is a questionable procedure if evoked sets vary across respondents. Urban [62] has proposed an extension to the regression approach by grouping respondents and regressing across individuals and choice alternatives. This provides many degrees of freedom (number of individuals times average evoked set) and allows for individuals with differing evoked sets; but it assumes the group is homogeneous with respect to the importance parameters \((2, i)\). Grouping procedures that identify homogeneity will be more fully discussed in the segmentation section of this paper.

This extended preference regression method was applied to the campus-wide HMO data base of 234 individuals. A saved data sample of 61 individuals was retained for predictive testing. The results are presented in Table V. The regression was done across individuals and stimuli and there were 642 total observations. The regression was significant at the 1\% level and all regression coefficients were significant at the 5\% level. The \(R^2\) fit statistic is not the most appropriate measure of fit since the dependent variable is rank ordered. Table V also reports a more appropriate measure—the fraction of times the predicted rank order preference was equal to the actual rank order preference. The observed fit of 0.44 can be compared to a fit of 0.25 obtained when all possible rank orderings are equally likely. The chi-square statistic for the matrix of actual versus predicted rank order was significant at the 5\% level. In addition, the fraction of times first preference was indicated was also reported as 0.48 (the random fraction is 0.25). These statistical results are quite encouraging and reflect adequate accuracy for compaction. Examination of the residuals indicated a unimodal distribution with 68\% of the observations between \(\pm 1\sigma\). The distribution was non-normal by the \(\chi^2\) test \([\chi^2 = 50.2 (df = 8)]\) and the departure was largely due to fluctuation in the tails. Use of a nonlinear form (logs of independent and dependent variables) and use of interaction terms did not improve the fits significantly.

Since errors were not normally distributed and the dependent observations in the regression were individual rank order preference judgments, the assumption of an interval scale required in linear regression may not be appropriate. To test this, a monotonic regression was conducted with Johnson's [26] monotonic regression program. This is essentially the same technique normally applied in conjoint analysis [17]. Although the fit is improved, the normalized importances are very similar to the linear
case. In the monotonic case, the importance of personalness is higher and of value is lower than in the linear case.

The estimated coefficients were used to predict the preference rank order for the 61 individuals in the saved data sample. The first preference was correctly predicted 36% of the time and the rank order 40%. The monotonic fit was again slightly better. The saved data fits are slightly lower than

<table>
<thead>
<tr>
<th>TABLE V</th>
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<tbody>
<tr>
<td>Preference Regressions</td>
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<thead>
<tr>
<th></th>
<th>Linear Regression</th>
<th>Monotonic Regression</th>
</tr>
</thead>
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<tr>
<td><strong>Coefficients (t statistic):</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Quality</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Raw</td>
<td>6.17 (9.16)</td>
<td>0.635</td>
</tr>
<tr>
<td>Normalized</td>
<td>0.32</td>
<td>0.33</td>
</tr>
<tr>
<td>Personalness</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Raw</td>
<td>3.86 (6.4)</td>
<td>0.495</td>
</tr>
<tr>
<td>Normalized</td>
<td>0.20</td>
<td>0.26</td>
</tr>
<tr>
<td>Value</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Raw</td>
<td>5.69 (12.2)</td>
<td>0.494</td>
</tr>
<tr>
<td>Normalized</td>
<td>0.30</td>
<td>0.26</td>
</tr>
<tr>
<td>Convenience</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Raw</td>
<td>3.34 (6.0)</td>
<td>0.326</td>
</tr>
<tr>
<td>Normalized</td>
<td>0.18</td>
<td>0.17</td>
</tr>
<tr>
<td><strong>Fit measures:</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>First Preference</td>
<td>0.48</td>
<td>0.52</td>
</tr>
<tr>
<td>Rank order</td>
<td>0.44</td>
<td>0.47</td>
</tr>
<tr>
<td>$R^2$</td>
<td>0.27</td>
<td>--</td>
</tr>
<tr>
<td>$\theta$</td>
<td>--</td>
<td>0.36</td>
</tr>
<tr>
<td>$\tau$</td>
<td>--</td>
<td>0.38</td>
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<tr>
<td><strong>Saved data fit:</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>First Preference</td>
<td>0.36</td>
<td>0.40</td>
</tr>
<tr>
<td>Rank order</td>
<td>0.40</td>
<td>0.42</td>
</tr>
</tbody>
</table>

the estimated results, as expected, but the drop is not alarming and the empirical results are supported by saved data testing.

While the statistical approach to compaction is encouraging, there are disadvantages to this approach. To gain degrees of freedom consumers must be grouped prior to estimation. Furthermore, no axiomatic theory is used to specify the functional form of compaction function. The degrees of freedom problem can be overcome by expanding the evoked set with abstract alternatives specified only by their attributes. On the basis of axioms that specify whether a function is measurable on a data set [59], conjoint analysis estimates individual compaction functions and derives importance weights by having consumers rank order factorially generated sets of these abstract alternatives. Although measurement theory allows
for polynomial combinations of nonlinear functions of the attributes, most
marketing applications have been limited to additive or multiplicative
combinations of these functions [17, 25], or simple interactions (Green
and Devita [15]). In these cases the attributes have been instrumental (e.g.,
price, brand name, package design) so that factorial combinations can be
unambiguously defined and design trade-offs can be explicitly made with
respect to controllable variables. Conjoint analysis was not applied to the
HMO case since at the early design stage interest is on psychological di-
mensons rather than instrumental variables. A study is now being con-
ducted to link instrumental variables such as HMO name, building, and
waiting time to the perceived quality. For an example of conjoint analysis
applied to health and hospital choice, see Wind and Spitz [67].

Since the conjoint axioms deal with measurement, they do not specify
what the functional form of the compaction function should be. On the
other hand, von Neumann-Morgenstern utility analysis draws on deductive
timey to derive unique functional forms from fundamental axioms and
verifiable assumptions describing how consumers evaluate alternatives.
These forms are important because they allow direct measurement of
risk aversion, importances, and interactions. However, until now utility
theory has been used exclusively for prescriptive decision making. Com-
paction adds the requirement that the theory allow stochastic choice,
i.e., that the probability of choosing a utility-maximizing alternative is
not necessarily certain. To use utility theoretic results for descriptive
choice, a theoretical construct of stochastic choice preference was defined
(a1 preferred to a2 means the probability of choosing a1 is greater than the
probability of choosing a2) and it was found necessary to augment the von
Neumann-Morgenstern axioms with a psychological choice axiom (Hauser
[22]), which can be shown to be similar to simple scalability (Luce [39],
Krantz [36], Tversky [60]). These additions enabled the establishment of
an isomorphism between utility and compaction theory and thus the ability
to apply many utility-theoretic results, including identification of func-
tional forms and direct assessment, to descriptive compaction theory. The
proofs are contained in [22].

Once the functional form is known, measurement proceeds as follows.
Rather than ranking factorially generated abstract alternatives, the con-
sumer is asked to consider abstract alternatives two at a time. One alter-
native is completely specified; the other leaves one characteristic (an at-
tribute level or an indicator of uncertainty) unspecified. The consumer’s
task is to specify that characteristic so that he is indifferent between the
two alternatives. If we assume that indifference means the consumer is
equally likely to choose either abstract alternative, then the compaction
values of the two alternatives are equal. Since the functional forms are
known, this provides one equation in the parameters. With sufficient in-
difference questions we can algebraically solve for the parameters. Calculations are based on the fact that \( x_{ij} \) indifferent to \( x_{i'} \) means \( c(x_{ij}, \lambda_i) = c(x_{i'}i, \lambda_i) \). (For examples of these calculations see [32].)

As a preliminary test of this technique, compaction functions over the four performance measures describing health-care delivery were assessed by a personal interview for a random sample of 80 members of the MIT student population. The compaction function was approximated with a relatively simple functional form, which was separable, multiplicative, and constantly risk averse in each dimension. Mathematically, the function, with individual specific parameters, is stated here for four performance measures.

\[
c_{ij} = \sum_m k_{im}u_{im}(x_{ijm}) + \sum_m \sum_{l>m} K_{li}k_{il}k_{in}u_{im}(x_{ijm})u_{il}(x_{ijl}) + \cdots + K_{ij}k_{ij}k_{in}u_{ii}(x_{ij}) \cdots u_{in}(x_{ijn})
\]

with \( u_{im}(x_{ijm}) = \frac{1 - \exp \left( -r_{im}x_{ijm} \right)}{1 - \exp \left( -r_{im}x_{m}^* \right)} \), and where

- \( c_{ij} \) = individual \( i \)'s scalar measure of goodness for alternative \( a_j \), i.e., the value of the compaction function, \( c_j(x_{ij}, \lambda_i) \) evaluated for individual \( i \) and alternative \( a_j \).
- \( u_{im}(x_{ijm}) \) = uni-attributed conditional "utility" scaling function. The form shown here is for constant risk aversion and is scaled from \( u_{im}(0) = 0 \) to \( u_{im}(x_{m}^*) = 1 \). (If \( r_{im} \to 0 \), \( u_{im}(\cdot) \) becomes the linear form used in statistical compaction.)
- \( x_{ijm} \) = the level of the \( m \)th performance measure as perceived by individual \( i \) for alternative \( a_j \).
- \( x_{m}^* \) = the maximum value of the \( m \)th performance measure.
- \( r_{im} \) = individual \( i \)'s risk-aversion coefficient relative to the \( m \)th performance measure.
- \( k_{im} \) = individual \( i \)'s importance coefficient for the \( m \)th performance measure.
- \( K_i \) = individual \( i \)'s interaction coefficient relative to the four performance measures.

Before assessment, the multiplicative form was selected on the basis of prior theory and in-depth interviews, aided by an interactive utility assessment computer program developed by Sicherman [52]. In a full-scale assessment, independence questions were used to check the validity of the assumptions necessary for the multiplicative form. It was found that these assumptions were correct for 66% of the respondents. Since this is the first time these assumptions have been tested on a consumer population, this result is encouraging. First we note that these assumptions are implicit in the functional forms used in conjoint analysis. Second we note that there are functional forms [7, 10, 29] that relax these assumptions, but as
yet it is not feasible to construct simple consumer measurements for these forms. Based on the results of administering the utility questionnaire (see "measurement" section of this methodology), the parameters \( \lambda_i = \{k_{i1}, k_{i2}, k_{i3}, r_{i1}, r_{i2}, r_{i3}, r_{i4}, K_i \} \) were calculated and are shown in Table VI. An empirical result unexpected in utility theory is that the risk-aversion coefficients and importances are highly correlated. This indicates that the student group is more concerned with risk for the more important performance measures. As a first comparison against the statistical technique, the individual specific perceptions, \( x_{ij} \), and preference parameters, \( \lambda_i \), were used to calculate scalar measures of goodness, \( c_{ij} \), for each alternative for each individual. When compared against rank-order preference, these resulted in a rank-order fit of 0.474 and a first preference fit of 0.495. These are in the same range as the fits of statistical procedure applied to the overall sample. Because of the differences between the statistical and direct compaction techniques, especially in their relation to the rest of the methodology, because of the nonlinear relationship between the factor scores and the directly measured performances, and because of the risk-averse scaling function, stronger and more explicit comparison tests need to be devised before importances can be compared. This is the subject of future research [24].

In summary, compaction identifies how consumers use the performance measures to evaluate alternatives. The key idea is that the consumer compacts the performance measures into a scalar measure of goodness for each alternative, which then compared across alternatives yields preference among the alternatives.

**Segmentation**

Development of a single product or service may not be the best strategy to exploit a potential market. For example, some people prefer high-quality service and are willing to pay the price, while others want adequate service at a low price. It may be that average service at an average price satisfies neither of these two groups. Thus, in the design of new products and services it is necessary to determine whether everyone has the same preferences relative to the reduced performance measures, or whether there exist segments of the population that have significantly different preferences. In the models of this methodology, this type of segmentation is represented by significantly different importance weights between segments, i.e., significantly different \( \Lambda_i \)'s.

This is a very attractive segmentation method since these parameters contain the key information used by the design team to determine managerial trade-offs in the attributes of the new product or service. This type of segmentation has been conceptually proposed by Haley [20] and is called "benefit" segmentation. The proposed segmentation is in contrast
<table>
<thead>
<tr>
<th></th>
<th>Importances ($k_{im}$)</th>
<th>Risk-Aversion Coefficients ($r_{im}$)</th>
<th>Interaction Coefficients ($K_i$)</th>
</tr>
</thead>
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<tr>
<td></td>
<td>Quality</td>
<td>Personal</td>
<td>Value</td>
</tr>
<tr>
<td>Median</td>
<td>0.800</td>
<td>0.525</td>
<td>0.706</td>
</tr>
<tr>
<td>Interquartile interval</td>
<td>0.650-</td>
<td>0.360-</td>
<td>0.492-</td>
</tr>
<tr>
<td></td>
<td>0.900</td>
<td>0.726</td>
<td>0.800</td>
</tr>
<tr>
<td>Interquartile range</td>
<td>0.250</td>
<td>0.367</td>
<td>0.308</td>
</tr>
</tbody>
</table>
to “life cycle” and “psychographics” as exemplified by Wells and Gruber [65]. See Frank, Massy, and Wind [12] for a summary of the literature on segmentation. Although these methods may be useful in creating ads for established products, we feel benefit segmentation is the best method for making attribute trade-offs inherent in new product design. After benefit segments are defined, they can be described by a profile of demographic and psychometric identifiers. The relative importance of these descriptions in defining segment membership can be examined by discriminant analysis [40].

In addition to the need for segmentation from a marketing strategy point of view, homogeneous groups are necessary for valid compaction if the preference regression technique is used. In the statistical estimation of compaction functions a fundamental paradox exists. It is theoretically sound to estimate importance weights only if the population is homogeneous with respect to these weights, but we do not have enough degrees of freedom for estimation unless we group individuals together. Therefore, we must segment the market before statistical estimation.

The description of segments has often been highly judgmental. Techniques such as clustering [14] and AID (Sonquist, Baker and Morgan [54]) have been used, but they are highly manipulative techniques and potentially misleading. See Doyle and Fenwick [6] for a discussion of AID dangers. We believe that these techniques can be used in the search for alternate segments, but that specific statistical tests are needed to determine the validity of the segmentation. We propose a statistical test that requires (1) the correlations to preference to be higher in each segment and (2) the importance coefficients to be significantly different across segments.

In the HMO case, cluster analysis and AID analysis techniques were used to search for segments. AID was used with the measured intent to join the MIT plan as the dependent variable and individual descriptors and demographics as the explanatory variables (Greer and Suuberg [18]). There were 690 observations, somewhat less than the recommended AID sample of 1000. The most significant segmentation variables were measures of an individual's current pattern of care (e.g., MIT health department versus private doctor for physicals and continuing care). However, the evidence was not very strong. The AID analysis explained 24% of the variation, while simulations based on random data explained 16%. Other attempts at segmentation were equally unconvincing. One attractive method of segmentation is to cluster the individual parameters of the utility function directly (see Table VI), but since the sample size for the utility study was only 80 students, clustering could not be used. This technique deserves future research effort.

Throughout the search for segmentation there was a common, but weak
<table>
<thead>
<tr>
<th></th>
<th>$R^2$</th>
<th>$F(df)$</th>
<th>Quality($t$)</th>
<th>Personalness($t$)</th>
<th>Value($t$)</th>
<th>Convenience($t$)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Overall (n=210)</td>
<td>0.27</td>
<td>80.0 (4,839)</td>
<td>6.17 (9.16)</td>
<td>3.86 (6.4)</td>
<td>5.69 (12.2)</td>
<td>3.34 (6.0)</td>
</tr>
<tr>
<td><strong>Partition One</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Faculty (n=67)</td>
<td>0.38</td>
<td>41.2 (4,263)</td>
<td>6.04 (5.61)</td>
<td>5.93 (5.61)</td>
<td>5.95 (9.60)</td>
<td>3.34 (3.5)</td>
</tr>
<tr>
<td>Student (n=80)</td>
<td>0.31</td>
<td>36.3 (4,315)</td>
<td>6.59 (6.10)</td>
<td>1.69 (1.87)</td>
<td>7.82 (8.26)</td>
<td>4.38 (4.38)</td>
</tr>
<tr>
<td>Staff (n=63)</td>
<td>0.19</td>
<td>14.3 (4,247)</td>
<td>4.98 (3.60)</td>
<td>5.15 (4.32)</td>
<td>3.54 (3.48)</td>
<td>2.08 (2.06)</td>
</tr>
<tr>
<td><strong>Partition Two</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Private (n=88)</td>
<td>0.28</td>
<td>34.0 (4,347)</td>
<td>6.05 (5.66)</td>
<td>4.60 (5.3)</td>
<td>4.86 (7.2)</td>
<td>4.53 (5.41)</td>
</tr>
<tr>
<td>MIT (n=109)</td>
<td>0.29</td>
<td>44.0 (4,431)</td>
<td>6.89 (6.91)</td>
<td>3.54 (3.76)</td>
<td>6.5 (3.76)</td>
<td>2.79 (3.57)</td>
</tr>
<tr>
<td>Mixed (n=12)</td>
<td>0.26</td>
<td>3.9 (4,43)</td>
<td>2.59 (1.0)</td>
<td>5.49 (2.0)</td>
<td>4.14 (1.62)</td>
<td>-1.03 (0.41)</td>
</tr>
</tbody>
</table>
indication that the pattern of existing care—MIT versus private—was a possible variable for segmentation. The “priors” of the health department were that the segmentation of faculty, students, and staff was important since these were operational segments, and it was felt that preferences varied across these groups.

Two alternative partitions were proposed and tested. The first was the prior segmentation of faculty-students-staff while the second was based on the multivariate analysis and was private versus MIT as an existing care supplier. Table VII presents the statistical regression across individuals and stimuli of preferences against the perceived attributes. If a segmentation is real, we require that (1) the regression fits be significantly better statistically in each segment than in the overall regression and (2) the importance coefficients be significantly different across segments. An informal examination of Table VII does not show much evidence of segmentation. The $R^2$ and $F$'s are reported on the basis of the finding that monotonic regression had little effect on the coefficients (see Table 5). The $R^2$ and $t$'s are significant but are not uniformly better within the segments. Only the faculty segment $R^2$ is improved substantially ($R^2 = 0.27$ overall versus 0.38 for faculty), but in this partition the $R^2$ for staff decreases substantially ($R^2 = 0.18$ versus 0.27 overall). Some differences in importances are evident between faculty and students with respect to personalness ($\lambda$ for faculty = 5.95 versus 1.69 for students). The $R^2$ in the second partition (private, MIT, or MIT and private existing care patterns) are almost the same as the overall value.

A formal statistical test has been described by Johnston [27] and Fisher [11] to test the significance of the difference of regressions. Assuming interval data and applying Johnston’s test, none of the segments in the partitions are significantly different at the 10% level ($F_{10}^{195} = 0.77$ for faculty/student/staff partition and $F_{10}^{195} = 0.50$ for the private/MIT/mixed pattern of care partition).

Thus, in the HMO data there is no statistical basis for benefit segmentation. This emphasizes the need for a formal statistical technique to test the segmentation identified by “priors,” the interpretive clustering, or AID techniques.

**Probability of Choice**

Perception, compaction, and segmentation provide diagnostics necessary to design successful new products and services, but to evaluate potential designs we must focus on estimating how many consumers will choose the new product or service and how many of the potential consumers will select each of the various competing alternatives. The probability of choice module of the methodology explicitly models how the consumer compares the scalar measures of goodness computed in the compaction
module. If we are far enough along in the design of a new product or service to have advertising copy and a sample product or pilot service available, the probability of choice model is estimated by observing how consumers choose among actual alternatives. Otherwise, as is the case early in the design process, actual new products or services are not yet available. In this case scenarios are given to approximate choice as closely as is feasible. For example, on the first cycle health services may be represented by concept statements, on the second cycle we may have finished brochures and video tape testimonials, and on the third a pilot program. Thus, when choice is possible we observe it, but when it is not possible we observe choice on proxy alternatives and augment the data with measures of intent.

Although the emphasis in early design of the HMO is on the discovery of perceptual dimensions, importances, and segments, a preliminary sales forecast is needed to see if the venture is viable. In the HMO case observed preference and intent measures are used to estimate actual choices.

Intent to enroll for MIT was measured on a 1–5 point scale (definitely yes, probably yes, might or might not, probably not, and definitely not). These intent measures were processed on the assumption that all the definite intent, half of the probable intent, and 30% of the might intent could be converted to choice at MIT. These are a little higher than those found by Gruber [19] for consumer brands (definite 0.755, probable 0.314, might 0.268), but we felt the MIT HMO marketing would result in more conversion and were consistent with Urban’s experience in test marketing [61]. The application of intent conversions resulted in an estimate that 23.3% of those aware of the HMO would enroll.

For consistency, the observed preference was converted to choice on the assumption that 80% of the first preference and 20% of the second preference would result in choice. These fractions are based on translations of preference to choice in simulated shopping studies of consumer products [53]. On the basis of these proportions it was found that 21.7% of those aware would enroll in the MIT HMO. Since this was consistent with the intent transformation value of 23.3%, a value of 23% was used in forecasting for the managerial decisions reported in a following section. In simulating the effects of a new design, the compaction coefficients (see Table V) were used along with new perceptual attribute values to predict individual preference rank order. These were converted to choice by assuming 80% of first preference and 20% of second preference would choose MIT if they were aware.

In late stages of design actual choice can be observed. In this case choices can be considered as the outcomes of Bernoulli probabilities, \( (p_s(a_i|c_a, \ldots, c_D)) \). These estimate each individual’s selection probabilities within a segment \( s \) for each alternative \( a_i \) conditioned on the
scalar goodness measures for each alternative \((c_{i1}, \cdots, c_{i\Omega})\). In cases where repetitive choice decisions are made by a consumer, separate trial and repeat choice parameters would be estimated based on the goodness measure before use and after use of the new product or service.

One approach is to use the multinomial logit model [42]. This model postulates that there is a "true utility," \(u_{ij}\), that completely describes a consumer's choice process (i.e., a consumer chooses the alternative \(a_i\) with the largest \(u_{ii}\)), but we can only observe part of \(u_{ij}\). In fact, the "true utility" is equal to an observable part, \(c_{ij}\), plus a random error. Deductive reasoning from distributional assumptions on the error term yields the multinomial logit model: 

\[
p(a_j|c_{i1}, \cdots, c_{i\Omega}) = \exp \{\beta c_{ij}\} / \sum_{j-1} \exp \{\beta c_{ij}\}.
\]

There are two ways to use this model in the methodology. The first is to substitute the compaction values directly into the logit model and estimate the parameter \(\beta\) by maximum likelihood. The second method is to specify the compaction function as a linear or nonlinear function of attributes and estimate parameters \((\lambda_j)\) by maximum likelihood. In this case \(\beta\) in the logit model is incorporated in the \(\lambda_i\). If \(c_i(x_{ij}, \lambda_i)\) is linear in the parameters, a number of programs are readily available [42].

The \(\beta\)-logit model has been calibrated and tested on data collected prior to national introduction of a new aerosol deodorant. The scalar measures of goodness were constant sum paired comparison scale values and the choice was the respondent's last purchase [53]. In this example, \(\beta\) was 2.09 with a \(t\) statistic of 10 (\(df=278\)). The goodness of fit of the model was evaluated based on a new set of entropy tests for a probability of choice model.

**Testing Probability of Choice Models**

Whether done on proxy choice and intent measures or on actual choice with the multinomial logit, the output of the probability of choice module is a predicted probability for each individual consumer for each choice alternative. In compaction we tested alternative models based on their ability to recover either first preference or rank-order preference; in probability of choice we would like another measure. We would like to discriminate between a model that assigns 0.9999 to the chosen alternative over one that assigns 0.5001 to an alternative. Traditional measures break down for this test because our predicted measures, probabilities, are scaled between 0 and 1 while our observed measures, choices, are nominally scaled (0 or 1). In fact, for this type of data it can be shown (Hauser [21]) that the expected value of tests such as least squares is optimized for values other than true probabilities. Instead we propose a test described in [21] based on honest reward theory [46] and information theory (Gallagher [13]). This test begins with a naturally occurring measure of uncertainty,
entropy, and computes the percent of that uncertainty that is explained by the probability model.

Specifically, if our null hypothesis is that all individuals have the same probability of choosing \( a_j \), then the entropy, \( H_0(A) \), relative to \( N_0 \) is given by \( H_0(A) = - \sum_{j=1}^{J} m_j^0 \log m_j^0 \) where \( m_j^0 \) is the null probability of choosing \( a_j \). The observed information, \( I(A; X) \), given by the model relative to the null hypothesis is given by: \( I(A; X) = \frac{1}{n} \sum_{i=1}^{n} \sum_{j=1}^{J} \delta_{ij} \log \left( \frac{\hat{p}_{ij}}{m_j} \right) \), where \( n \) is the total number of consumers, \( P_{ij} \) are the probabilities predicted by our model, and \( \delta_{ij} = 1 \) if \( i \) chooses \( j \) and \( \delta_{ij} = 0 \) otherwise. The percent of uncertainty, \( U^2 \), explained by the model relative to the null case is simply \( I(A; X)/H_0(A) \) and the residual uncertainty is simply \( 1 - U^2 \).

In the deodorant case, this information test was applied. The null hypothesis, \( N_0 \), was that the probability of choice was proportional to the empirical share of people who chose each alternative. The \( U^2 \) value was 74\% (\( H_0(A) = 2.23 \) and \( I(A; X) = 1.64 \)). See [21] and [53] for discussion and application of this test. This indicates substantial contribution of the model to reduction of uncertainty and, along with the \( t \) statistic for \( \beta \) of 10, suggests highly significant results. Probability of choice predicts individual choice for products or services in an individual’s evoked set; the next step aggregates the individual choice probabilities.

**Aggregation**

The final step in the consumer-response process model is aggregation. It combines the individual choice probabilities to produce numerical estimates of the total share of choices and number of people choosing each alternative. If relevant population segments were identified in segmentation, aggregation explicitly uses them to extrapolate from the sample population to the target population. In addition to expected choices, the variances on the group choice are also useful in considering the risk associated with the new venture.

In most applications, individual choice probabilities are roughly independent; thus the central limit theorem can be used to calculate the joint probability distribution of the market shares. The grand means are determined by averaging the individual results obtained by direct substitution of the individual goodness measures in the probability of choice models \( (P_{ij}) \). The variance is based on the sum of the individual variances \( (P_{ij}(1-P_{ij})) \). In some cases it is convenient to represent a distribution across the goodness measures, in which case one must integrate rather than sum to find the mean and variance (see Koppelman [35]). Explicitly the mean market share for alternative \( a_j \) is given by \( m_{s_j} = (1/n) \sum_{i=1}^{n} P_{ij} \).

The variance is given by \( \text{var} (m_{s_j}) = (1/n) \sum_{i=1}^{n} P_{ij}(1-P_{ij}) \).

The final step in aggregation reflects the fact that all the measurement
and estimation is based on consumers who evoked the alternative by past experience or who were made aware of the alternative in the interview. The adjustment is to multiply the aggregate share of choices for an alternative times the estimated level of evoking and awareness that the new product or service is expected to achieve on the basis of the predicted level of marketing effort or the evoking rates for competing alternatives.

In the HMO case, since the predicted probabilities were discrete, the summation technique was used. Evoking and awareness were estimated judgmentally on the basis of a planned marketing effort. The empirical results are discussed in the case application.

This completes the analytic discussion of the methodology. The primary purpose of the methodology is to aid in the design of new products and services. The next section illustrates this use by discussing the managerial implications of the HMO case.

4. CASE APPLICATION

The proposed methodology for modeling consumer response to innovation was applied to the problem of converting the MIT health department into an HMO. Some of the empirical findings have been cited earlier. This section will concentrate on the managerial use of the model in evaluation and refinement of the HMO design.

The model estimation was based on consumer interviews of 447 faculty, students, and staff. Of these, 367 were prospective members of the HMO and 80 were members of a pilot HMO begun a year earlier.

First perceptual maps were derived by a factor analysis of the ratings of existing and new health plan descriptions. Figure 6 gives the overall average factor scores (see Table I for a concept statement, Table II for the rating scales, and Table IV for factor loadings).

First, it should be noted that the average perceptions of the existing care system are better than those of the MIT HMO concept based on the measures from the prospective members on all dimensions except convenience. Next, note that the MIT HMO concept is perceived better than the Harvard Community Health Plan (HCHP) on all dimensions except quality. The examination of the factor score coefficients and raw ratings (Figure 5) showed the lower quality rating for the MIT HMO was based almost entirely on a low score for MIT on hospital quality. This was because the MIT plan hospital (i.e., Cambridge City Hospital) was smaller and not as well regarded as the prestigious Boston hospitals. The Harvard Community Plan was higher on quality because of its hospital ratings (Beth Israel, Peter Bent Brigham, and Childrens Hospital).

When comparing existing care perceptions with the perceptions of those in the MIT pilot plan, a substantially different picture emerges. The MIT pilot plan exceeds the average existing care perceptions on all
dimensions. The stated intent of over 90% to re-enroll supported the notion that the plan was very effective as seen by those in the pilot plan. It is clear that perceptions based on actual pilot plan performance of the HMO are much better than perceptions of the plan based on the concept statement. Although this could be post-purchase rationalization, it is more likely that this is a case of a good product where few people perceive it as such until using it.

Figure 6. Overall average factor scores. MIT Pilot = O—O; MIT HMO Concept = ⃝—��; Existing = X—X; Harvard Community Plan = □—□.

The perceptual maps indicated two major managerial findings. First, "quality" of the plan was low and probably would be improved by better hospital affiliation. Second, if the HMO was to be successful, it would have to develop an aggressive campaign to communicate actual plan performance to perspective members.

Analysis of the preferences indicated that overall the rank order of importance of the four relevant attributes was quality, value, personal-ness, and convenience.

The choice model and empirical observations were used to forecast new enrollment and re-enrollment for the next year. Table VIII gives the
forecast based on the estimated probability of enrollment, given that the respondent was aware at the level presented by the survey and based upon an estimate of how many potential users would be aware at this level. Since the pilot HMO was made available only to faculty and staff, there

**TABLE VIII**

**Forecast of MIT Enrollment**

*Case I. Existing Design*

<table>
<thead>
<tr>
<th>Groups</th>
<th>No. Not Now in Pilot HMO</th>
<th>Enrollment if Aware</th>
<th>Estimated Awareness</th>
<th>Estimated Enrollment</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>New enrollment:</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Students</td>
<td>8000</td>
<td>× 33%</td>
<td>× 70%</td>
<td>= 1,848</td>
</tr>
<tr>
<td>Faculty</td>
<td>3800</td>
<td>× 15%</td>
<td>× 70%</td>
<td>= 399</td>
</tr>
<tr>
<td>Staff</td>
<td>3400</td>
<td>× 22%</td>
<td>× 70%</td>
<td>= 523</td>
</tr>
<tr>
<td><strong>Overall</strong></td>
<td>17,200</td>
<td>23%</td>
<td>70%</td>
<td>2,770</td>
</tr>
<tr>
<td><strong>Re-enrollment:</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Existing HMO</td>
<td>Repeat Rate</td>
<td>Estimated to Remain at MIT</td>
<td></td>
<td></td>
</tr>
<tr>
<td>subscribers</td>
<td>1067</td>
<td>× 92.5%</td>
<td>× 86.3%</td>
<td>= 852</td>
</tr>
<tr>
<td><strong>Total enrollment</strong></td>
<td>= 3622</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

*Case II. Improved Design*

<table>
<thead>
<tr>
<th>Groups</th>
<th>No. Not Now in Pilot Program</th>
<th>Enrollment if Aware</th>
<th>Estimated Awareness</th>
<th>Estimated Enrollment</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>New enrollment:</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Students</td>
<td>8000</td>
<td>× 42%</td>
<td>× 85%</td>
<td>= 2,856</td>
</tr>
<tr>
<td>Faculty</td>
<td>3800</td>
<td>× 25%</td>
<td>× 85%</td>
<td>= 808</td>
</tr>
<tr>
<td>Staff</td>
<td>3400</td>
<td>× 30%</td>
<td>× 85%</td>
<td>= 867</td>
</tr>
<tr>
<td><strong>Overall</strong></td>
<td>17,200</td>
<td>31%</td>
<td>85%</td>
<td>4,331</td>
</tr>
<tr>
<td><strong>Re-enrollment:</strong></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Existing HMO</td>
<td>Repeat Rate</td>
<td>Estimated to Remain at MIT</td>
<td></td>
<td></td>
</tr>
<tr>
<td>subscribers</td>
<td>1067</td>
<td>× 95%</td>
<td>× 86.3%</td>
<td>= 874</td>
</tr>
<tr>
<td><strong>Total enrollment</strong></td>
<td>= 5,405</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

was interest in how many additional faculty and staff would enroll. Therefore, the total new enrollment shown in Table VIII is decomposed into three groups. Re-enrollment is based on the empirically estimated probability of re-enrollment (92.5%) and estimates of migration out of the MIT community. The total enrollment forecast of 3600 families was just financially sufficient to maintain the HMO. Considering the inherent risk involved in any new service venture, the decision to expand the existing pilot HMO could not be supported on the basis of the initial design.
However, the existing pilot was not the best design. The use of aggressive communication to close the gap of perceptions and performance and the change of hospital associations were identified as methods of improvement. The model was used to simulate the effects of these design changes by assuming (1) the communication campaign and the association with prestigious hospitals could move the perceptions one half of the distance from concept to actual pilot performance on quality, personalness, and value (see Figure 6), and (2) 85% awareness would be created rather than 70%. The consumer response model forecast an enrollment of 5400 families on the basis of estimates of importance weights and a linear compaction function produced by the regression of factor scores against preference (see Table V), and the conversion of rank order preference to choice described in the compaction section of this paper. Other simulations were based on assumptions of competition from the Harvard Community Health Plan (HCHP), which indicated the enrollment could drop to 2400 with HCHP offered at MIT and no improved design. The most likely forecast was based on HCHP’s being offered with an improved MIT plan and was for 4950 family enrollments. This was not sufficient to make a positive recommendation for full implementation and a new building, even when based on the response of consumers to the revised communication and design strategy. The predictions did substantiate the demand for more HMO services and supported a plan to expand within the existing physical facilities to 3000 faculty and staff families. During this expansion the hospital affiliation was shifted from Cambridge City Hospital to Mount Auburn Hospital with referral to Massachusetts General Hospital. MIT is now expanding its HMO to meet the indicated need and will consider marketing its HMO to students as facilities become available.

From this initial application it appears that the consumer response model is relevant to the management of innovation and can be useful in improving designs of new products or services, forecasting the acceptance of innovations, and reducing the risk of failure.

5. SUMMARY AND FINAL COMMENTS

This paper proposes and presents evidence for a normative methodology to elicit and guide creativity in the design of innovative products and services. The strength of the methodology is that it effectively integrates state-of-the-art analysis techniques from the fields of psychometrics, utility theory, and stochastic choice theory in a scheme that is oriented toward the needs and desires of managers and staff responsible for innovation. Its primary use is to enhance early creative identification and design of high potential products and services by providing important diagnostics that describe consumers’ perceptions of the alternatives and consumers’ preferences relative to measures of these perceptions. It also shows how to
identify and test managerially relevant segments based on homogeneity of preference and gives numerical indications of consumer response within each segment. Diagnostics are produced for design insight, but the methodology also simulates and evaluates quantitative and qualitative design changes in the characteristics of the alternatives or in the implementation strategy. Thus managers can readily test and improve the intuition they develop on the basis of previous experience and the previous outputs of the methodology.

Our new techniques will be tested in future work. Attention will be directed at confirming the validity of (1) the new measurement instruments that allow mass direct assessment of consumers’ utility functions, (2) the use of psychometrics to get a complete and parsimonious set of performance measures for utility assessment, (3) the criteria and statistical test of segmentation based on preference parameters, (4) the stochastic interpretation and choice axiomization that yields an isomorphism between utility-theoretic results and compaction, and (5) the information-theoretic testing techniques for assessing the accuracy of predicted individual probabilities.

In addition to testing the new techniques, we need comparative research on the alternative methods of compaction. The techniques of expectancy value, preference regression, conjoint analysis, and utility theory should be analyzed to determine theoretically and empirically when each one is most appropriate. Some work is under way [23, 24]. Utility theory needs new measurement methods that are less demanding on the consumer and can be executed in the relatively short time available in a personal interview. In addition, the empirical results from applying utility theory in the HMO case indicate error in fit and predictions, yet utility theory has no structure to deal with measurement error. This is a high priority area for research. The multinomial logit model proposed for the modeling of probability of choice is basically a zero-order model. The probability of purchases does not depend on past purchasing except through preference values. In many situations switching and purchase loyalty are phenomena to consider (see [41]). Research is needed in the formulation of higher order stochastic models that include preference and switching phenomena. A final research area is in the complementary use of conjoint analysis and perceptual mapping. Research is needed to develop procedures so that conjoint analysis can be used to link the instrumental variables of the product to the perceptual performance dimension.

The methodology has been applied to the repositioning of the Master of Science program at MIT’s Sloan School of Management, the design of financial service packages (banking and brokerage), and the positioning of new frequently purchased consumer products (e.g., antacids, personal care products, and pain relievers). In 10 applications managerial reception
has been enthusiastic and several high-potential products have been designed. Applications are now planned to develop consumer acceptance of squid as a protein source, to design telecommunications technology, and to design both suburban and intercity passenger transportation systems. Although the initial experience is encouraging future application, testing and research will be needed to fully document and develop the impact of applying the proposed methodology to the design of new products and services.

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